

Lösningar till tentamen i Matematisk statistik och diskret matematik D2 (MVE055/MSG810).

Den 28 augusti 2013. These are sketches of the solutions.

1. Lösning:
Write

$$\begin{aligned}\mathbb{P}(U = 1, V = 1) &= \frac{1}{6}, & \mathbb{P}(U = -1, V = 1) &= \frac{1}{3} \\ \mathbb{P}(U = 1, V = -1) &= \frac{1}{3}, & \mathbb{P}(U = -1, V = -1) &= \frac{1}{6}\end{aligned}$$

- a) $x^2 + Ux + V$ has a real root iff $U^2 - 4V \geq 0$ which means $V = -1$. Clearly, if $V = -1$, then $U^2 - 4V = 5$. So the probability of a real root is $\frac{1}{2}$.
- b) The expected value of the larger root is

$$\begin{aligned}& \frac{(-1 + \sqrt{5})}{2} \mathbb{P}(U = 1|V = -1) + \frac{(1 + \sqrt{5})}{2} \mathbb{P}(U = -1|V = -1) \\ &= \frac{1}{(\mathbb{P}(V = -1))} \left(\frac{(-1 + \sqrt{5})}{2} \frac{1}{3} + \frac{(1 + \sqrt{5})}{2} \frac{1}{6} \right) = \frac{\sqrt{5}}{2} - \frac{1}{6}.\end{aligned}$$

- c) $x^2 + Wx + W$ has a real root if $W^2 - 4W \geq 0$. If $W = U + V$, W takes values 2, 0, -2, and the equation has a real root of $W = 0$ or -2 . Then $\mathbb{P}(W = 0) = \frac{2}{3}$ and $\mathbb{P}(W = 0 \text{ or } -2) = \frac{5}{6}$.

Lösning:

2. a) The sample mean \bar{X} has mean μ and variance $\frac{\sigma^2}{n}$. Hence, by Chebyshev's inequality

$$\mathbb{P}(|\bar{X} - \mu| \geq 2\sigma) \leq \frac{\sigma^2}{n(2\sigma)^2} = \frac{1}{4n}.$$

Thus $n = 25$ is sufficient. If more is known about the distribution of X_i , then a smaller sample size may suffice.

- b) If $X_i \sim N(\mu, \sigma^2)$, then $\bar{X} - \mu \sim N(0, \frac{\sigma^2}{n})$, and probability $\mathbb{P}(|\bar{X} - \mu| \geq \sigma)$ equals

$$\mathbb{P}\left(\frac{|\bar{X} - \mu|}{\left(\frac{\sigma^2}{n}\right)^{\frac{1}{2}}} \geq \frac{\sigma}{\left(\frac{\sigma^2}{n}\right)^{\frac{1}{2}}}\right) = \mathbb{P}(|Z| \geq n^{\frac{1}{2}})$$

where $Z \sim N(0, 1)$. But $\mathbb{P}(|Z| \geq 2.58) = 0.99$, and so we require that $n^{\frac{1}{2}} \geq 2.58$, i.e. $n \geq 7$. As we see, knowledge that the distribution is normal allows a much smaller sample size, even to meet a tighter conditions.

3. Lösning:

- a) Label states by A, B, C indicating which player is not playing in a given game. Then the transition matrix is $\{A, B, C\} \times \{A, B, C\}$:

$$\begin{pmatrix} 0 & \frac{s_C}{(s_B+s_C)} & \frac{s_B}{(s_B+s_C)} \\ \frac{s_C}{(s_A+s_C)} & 0 & \frac{s_A}{(s_A+s_C)} \\ \frac{s_B}{(s_A+s_B)} & \frac{s_A}{(s_A+s_B)} & 0 \end{pmatrix}$$

The process is a Markov chain because the results of the subsequent games are independent.

- b) Here, we look for the probability that after three steps the chain returns to a given initial state.
From the symmetry, this probability is the same for any choice of the initial state and is equal to

$$p_{ABPBCPCA} + p_{ACPCBPBA} = \frac{2s_A s_B s_C}{(s_A + s_B)(s_B + s_C)(s_C + s_A)}$$

Lösning:

4. a) As an estimate of μ , we report the sample mean $\bar{T} = \frac{1}{10} \sum_{i=1}^{10} T_i = 1.80$. For a 90% confidence interval, we first obtain the percentile $z_{0.95} = 1.645$ and then compute the interval

$$\left[\bar{T} - z_{0.95} \sqrt{\frac{0.004}{10}}, \bar{T} + z_{0.95} \sqrt{\frac{0.004}{10}} \right] = [1.767, 1.833].$$

- b) We report $S^2 = \left(\frac{1}{9}\right) \sum_{i=1}^{10} (T_i - 1.80)^2 = 0.0051$ as an estimate of σ^2 . For a 90% confidence interval, we obtain percentiles

$$\chi_{9,0.95}^2 = 16.919$$

$$\chi_{9,0.05}^2 = 3.325.$$

We then compute the interval

$$\left[\frac{(n-1)S^2}{\chi_{9,0.95}^2}, \frac{(n-1)S^2}{\chi_{9,0.05}^2} \right] = [0.0027, 0.0138].$$