

## Solutions to Chapter 6 Exercises (Part 2)

### Problem 6.3

(a) Define

$$I_n(z) = \int \int \dots \int_{\sum_{i=1}^n x_i \leq z} dx_1 dx_2 \dots dx_n.$$

Then  $c = I_N^{-1}(1)$ . Note that

$$\begin{aligned} I_n(z) &= \int_0^z \left[ \int \int \dots \int_{\sum_{i=1}^{n-1} x_i \leq z - x_n} dx_1 dx_2 \dots dx_{n-1} \right] dx_n \\ &= \int_0^z I_{n-1}(z - x_n) dx_n \\ &= \int_0^z I_{n-1}(u) du. \end{aligned}$$

Using this iteration, we see that

$$I_1(z) = \int_0^z du = z,$$

$$\begin{aligned}
I_2(z) &= \int_0^z u du = \frac{1}{2}z^2, \\
I_3(z) &= \int_0^z \frac{1}{2}u^2 du = \frac{1}{3!}z^3, \\
&\vdots \\
I_n(z) &= \frac{1}{n!}z^n.
\end{aligned}$$

Then,  $c = I_N^{-1}(1) = N!$ .

(b)

$$\begin{aligned}
f_{X_1, X_2, \dots, X_M}(x_1, x_2, \dots, x_M) &= \int \int \dots \int_{x_{M+1} + \dots + x_N \leq 1 - x_1 - x_2 - \dots - x_M} c dx_N dx_{N-1} \dots dx_{M+1} \\
&= c I_{N-M}(1 - x_1 - x_2 - \dots - x_M) \\
&= \frac{N!}{(N-M)!} (1 - x_1 - \dots - x_M)^{N-M}, \quad \sum_{m=1}^M x_m \leq 1, x_i \geq 0
\end{aligned}$$

(c) From part(b), we have

$$f_{X_i}(x_i) = N(1 - x_i)^{N-1}.$$

Then,

$$f_{X_1}(x_1)f_{X_2}(x_2)\dots f_{X_N}(x_N) = N^N(1-x_1)^{N-1}\dots(1-x_N)^{N-1} \neq f_{X_1, X_2, \dots, X_N}(x_1, x_2, \dots, x_N) = N!$$

Thus,  $X_i$  are identically distributed, but they are not independent.

## Problem 6.12

$$\mathbf{X} = [X_1, X_2, X_3, X_4]^T$$

$$\boldsymbol{\Omega} = [\omega_1, \omega_2, \omega_3, \omega_4]^T$$

From problem 6.11 we know

$$\begin{aligned}
\mathbf{X} &= [X_1, X_2, X_3, X_4]^T \\
\Phi_{\mathbf{X}}(\boldsymbol{\Omega}) &= \exp\left(-\frac{\boldsymbol{\Omega}^T \mathbf{C}_{\mathbf{XX}} \boldsymbol{\Omega}}{2}\right) \\
\rightarrow E\left[e^{j(\omega_1 X_1 + \omega_2 X_2 + \omega_3 X_3 + \omega_4 X_4)}\right] &= \exp\left(-\frac{\boldsymbol{\Omega}^T \mathbf{C}_{\mathbf{XX}} \boldsymbol{\Omega}}{2}\right)
\end{aligned}$$

Expanding both the exponentials and equating the coefficients of  $\omega_1 \omega_2 \omega_3 \omega_4$  we get

$$\begin{aligned}
& E \left[ 1 + j(\omega_1 X_1 + \omega_2 X_2 + \omega_3 X_3 + \omega_4 X_4) - \frac{1}{2!} (\omega_1 X_1 + \omega_2 X_2 + \omega_3 X_3 + \omega_4 X_4)^2 + \dots \right] \\
& = \exp \left( 1 - \frac{\Omega^T \mathbf{C}_{\mathbf{X}\mathbf{X}} \Omega}{2} + \frac{1}{2!} \left( \frac{\Omega^T \mathbf{C}_{\mathbf{X}\mathbf{X}} \Omega}{2} \right)^2 + \dots \right) \\
& \rightarrow E [ \dots + X_1 X_2 X_3 X_4 \omega_1 \omega_2 \omega_3 \omega_4 + \dots ] = \left( \dots + \frac{1}{2!} \frac{1}{2^2} (c_{12} c_{34} + c_{13} c_{24} + c_{14} c_{23}) 8 \omega_1 \omega_2 \omega_3 \omega_4 + \dots \right)
\end{aligned}$$

Where  $c_{ij} = \text{cov}(X_i, X_j)$

$$\begin{aligned}
& \rightarrow E [ X_1 X_2 X_3 X_4 ] = \left( \frac{1}{2!} \frac{1}{2^2} (c_{12} c_{34} + c_{13} c_{24} + c_{14} c_{23}) 8 \right) \\
& \rightarrow E [ X_1 X_2 X_3 X_4 ] = (c_{12} c_{34} + c_{13} c_{24} + c_{14} c_{23})
\end{aligned}$$

Since  $X_i$  are zero mean random variables  $c_{ij} = E[X_i X_j]$

$$\rightarrow E [ X_1 X_2 X_3 X_4 ] = E [ X_1 X_2 ] E [ X_3 X_4 ] + E [ X_1 X_3 ] E [ X_2 X_4 ] + E [ X_1 X_4 ] E [ X_2 X_3 ]$$

## Problem 6.14

$$\begin{aligned}
\frac{d}{dy} F_{Y_m}(y) &= \frac{d}{dy} \sum_{k=m}^N \binom{N}{k} \left( F_X(y) \right)^k \left( 1 - F_X(y) \right)^{N-k} \\
&= \sum_{k=m}^N \binom{N}{k} k f_X(y) \left( F_X(y) \right)^{k-1} \left( 1 - F_X(y) \right)^{N-k} \\
&\quad - \sum_{k=m}^N \binom{N}{k} (N-k) f_X(y) \left( F_X(y) \right)^k \left( 1 - F_X(y) \right)^{N-k-1}
\end{aligned}$$

Since  $k=N$  evaluate the term  $(N-k)$  in the second sum to zero, we can set the range of  $k$  to be  $k=m, \dots, N-1$ .

$$\begin{aligned}\frac{d}{dy} F_{Y_m}(y) &= \sum_{k=m}^N \frac{N!}{(k-1)!(N-k)!} f_X(y) (F_X(y))^{k-1} (1-F_X(y))^{N-k} \\ &\quad - \sum_{k=m}^{N-1} \frac{N!}{(k)!(N-k-1)!} f_X(y) (F_X(y))^k (1-F_X(y))^{N-k-1}\end{aligned}$$

In the second sum, let  $n=k+1$ . Then

$$\begin{aligned}f_{Y_m}(y) &= \frac{d}{dy} F_{Y_m}(y) = \sum_{k=m}^N \frac{N!}{(k-1)!(N-k)!} f_X(y) (F_X(y))^{k-1} (1-F_X(y))^{N-k} \\ &\quad + \sum_{n=m+1}^N \frac{N!}{(n-1)!(N-n)!} f_X(y) (F_X(y))^{n-1} (1-F_X(y))^{N-n}\end{aligned}$$

The second series cancel the first series except the terms  $k=m$ . Hence

$$f_{Y_m}(y) = \frac{N!}{(m-1)!(N-m)!} f_X(y) (F_X(y))^{m-1} (1-F_X(y))^{N-m}$$

### Problem 6.17

(a)

$$E[Y_n] = E\left[\frac{X_n + X_{n-1}}{2}\right] = \frac{1}{2}E[X_n] + \frac{1}{2}E[X_{n-1}] = 0.$$

$$\begin{aligned}
E[Y_n Y_k] &= E \left[ \left( \frac{X_n + X_{n-1}}{2} \right) \left( \frac{X_k + X_{k-1}}{2} \right) \right] \\
&= \frac{1}{4} (E[X_n X_k] + E[X_n X_{k-1}] + E[X_{n-1} X_k] + E[X_{n-1} X_{k-1}]) \\
&= \frac{1}{4} [c_{n,k} + c_{n,k-1} + c_{n-1,k} + c_{n-1,k-1}]
\end{aligned}$$

Using  $c_{i,j} = \delta_{i,j}$  we get

$$E[Y_n Y_k] = \begin{cases} \frac{1}{2} & n = k, \\ \frac{1}{4} & n = k \pm 1, \\ 0 & \text{otherwise.} \end{cases}$$

The correlation matrix is then of the form

$$\mathbf{R} = \frac{1}{4} \begin{bmatrix} 2 & 1 & 0 & 0 & 0 & \dots \\ 1 & 2 & 1 & 0 & 0 & \dots \\ 0 & 1 & 2 & 1 & 0 & \dots \\ 0 & 0 & 1 & 2 & 1 & \dots \\ \vdots & \vdots & \vdots & \vdots & \vdots & \ddots \end{bmatrix}.$$

(b) In this case,  $c_{i,j} = \sigma_X^2 \delta_{i,j}$  so that

$$\mathbf{R} = \frac{\sigma_X^2}{4} \begin{bmatrix} 2 & 1 & 0 & 0 & 0 & \dots \\ 1 & 2 & 1 & 0 & 0 & \dots \\ 0 & 1 & 2 & 1 & 0 & \dots \\ 0 & 0 & 1 & 2 & 1 & \dots \\ \vdots & \vdots & \vdots & \vdots & \vdots & \ddots \end{bmatrix}.$$