SERIK SAGITOV, Chalmers and GU, February 14, 2018

## Solutions chapter 10

## Problem 10.2

IID sample  $(X_1, \ldots, X_n)$  from the uniform distribution U(0, 1).

(a) For a fixed x, the empirical distribution function  $F_n(x)$  is the sample proportion estimate of p = F(x) = x. The variance of  $F_n(x)$  is

$$\sigma_{F_n(x)}^2 = \frac{p(1-p)}{n} = \frac{x(1-x)}{n}$$

the that

$$\sigma_{F_n(x)} = \sqrt{\frac{x(1-x)}{n}}, \quad x \in [0,1].$$

(b) Generate 100 samples of size n = 16. For each sample plot  $F_n(x) - F(x)$  and relate what you see to your answer to (a). Matlab code

#### Problem 10.5

We have

$$F_n(u) = \frac{1}{n} [1_{\{X_1 \le u\}} + \ldots + 1_{\{X_n \le u\}}], \quad \mathcal{E}(F_n(u)) = F(u),$$
  
$$F_n(v) = \frac{1}{n} [1_{\{X_1 \le v\}} + \ldots + 1_{\{X_n \le v\}}], \quad \mathcal{E}(F_n(v)) = F(v).$$

Assuming u < v, we get

$$E(F_n(u) \cdot F_n(v)) = \frac{1}{n^2} \left[ \sum_{i=1}^n E(1_{\{X_i \le u\}} 1_{\{X_i \le v\}}) + \sum_{i=1}^n \sum_{j \ne i} E(1_{\{X_i \le u\}} 1_{\{X_j \le v\}}) \right]$$
$$= \frac{1}{n^2} \left[ \sum_{i=1}^n F(u) + \sum_{i=1}^n \sum_{j \ne i} F(u) F(v) \right]$$
$$= \frac{1}{n} [F(u) + (n-1)F(u)F(v)].$$

Finish by using

$$\operatorname{Cov}(F_n(u), F_n(v)) = \operatorname{E}(F_n(u) \cdot F_n(v)) - \operatorname{E}(F_n(u)) \cdot \operatorname{E}(F_n(v)).$$



# Problem 10.6

Ordered sample n = 59

25% quantile

50% quantile

75% quantile

(a) Use Matlab commands

x=data vector;	
stairs(sort(x),(1:length(x))/length(x))	% empirical cdf
hist(x)	% histogram, the same as hist(x,10)
normplot(x)	% normal probability plot
prctile(x,90)	% 0.90-quantile

The distribution appears to be rather close to normal. The 10% quantile

$$\frac{X_{(6)} + X_{(7)}}{2} = \frac{13.66 + 13.68}{2} = 13.67$$



(b) Since  $\bar{x} = 14.58$  and s = 0.78, the one-sided 99% of the population distribution for the natural wax is

$$(-\infty, 14.58 + 2.33 \cdot 0.78) = (-\infty, 16.40).$$

Expected means

 $\begin{array}{ll} 1\% \mbox{ dilution } & \mu_1 = 14.58 \cdot 0.99 + 85 \cdot 0.01 = 15.28 & \mbox{ can not be detected} \\ 3\% \mbox{ dilution } & \mu_3 = 14.58 \cdot 0.97 + 85 \cdot 0.03 = 16.69 & \mbox{ can be detected} \\ 5\% \mbox{ dilution } & \mu_5 = 14.58 \cdot 0.95 + 85 \cdot 0.05 = 18.10 & \mbox{ can be detected} \\ \end{array}$ 

## Problems 10.11, 10.13, 10.14.

$$1 - F(t) = e^{-\alpha t^{\beta}}, \quad f(t) = \alpha \beta t^{\beta - 1} e^{-\alpha t^{\beta}}, \quad h(t) = \alpha \beta t^{\beta - 1}.$$

- If  $\beta = 1$ , then  $h(t) = \alpha$  is constant and the distribution is memoryless.
- If  $0 < \beta < 1$ , then h(t) decreases with t meaning that the longer you live the healthier you become.
- If  $\beta > 1$ , then h(t) increases with t meaning that the older individuals die more often than the younger.

#### Problem 10.29

Stem and leaf display for n = 26 observations including k = 5 outliers:

133:7 134:134 135:002244 135:88 136:36 High: 141.2, 143.3, 146.5, 147.8, 148.8

Let N be the number of outliers in a non-parametric bootstrap sample.

(a) Due to sampling with replacement we have  $N \sim \text{Bin}(26, \frac{5}{26})$ .

(b) Find  $P(N \ge 10)$ :

$$P(N \le 9) = binocdf(9, 26, 5/26) = 0.9821,$$
  
 $P(N \ge 10) = 1 - 0.9821 = 0.018.$ 

(c) In B = 1000 bootstrap samples, we expect

$$B \cdot P(N \ge 10) = 18$$

samples to contain 10 or more of outliers.

(d) The probability that a bootstrap sample is composed entirely of these outliers is

$$P(N = 25) = (5/26)^{26} = 2.4 \cdot 10^{-19}.$$

## Problem 10.37

Same data as in Problem 10.6.

(a) The Matlab commands

trimmean(x,10)trimmean(x,20)

give  $\bar{X}_{0.1} = 14.586$  and  $\bar{X}_{0.2} = 14.605$ .

m = trimmean(X, percent) calculates the trimmed mean of the values in X. For a vector input, m is the mean of X, excluding the highest and lowest k data values, where  $k=n^*(percent/100)/2$  and where n is the number of values in X.

(b) An approximate 90% CI for the mean is

$$14.58 \pm 1.645 \cdot \frac{0.78}{\sqrt{59}} = 14.58 \pm 0.17 = (14.41; 14.75)$$

(c) Nonparametric 90% CI for the population median M is  $(X_{(k)}, X_{(60-k)})$ , where P(Y < k) = 0.05and  $Y \sim Bin$  (59, 0.5). Applying the normal approximation for Bin (n, p) with continuity correction

$$\mathbf{P}(Y < k) = \mathbf{P}(Y \le k - 1) \approx \Phi\left(\frac{k - 0.5 - np}{\sqrt{np(1 - p)}}\right),$$

we arrive at equation

$$\frac{k - 0.5 - \frac{59}{2}}{\sqrt{\frac{59}{4}}} = -1.645,$$

that gives k = 24. This yields

$$(X_{(k)}, X_{(60-k)}) = (X_{(24)}, X_{(36)}) = (14.43; 14.75).$$

(d) The Matlab commands for the non-parametric bootstrap

```
n=59; B=1000;

z=x(random('unid',n,n,B)); \% ('unid',n) - uniform discrete [1, n], 1000 samples of size n

t1=trimmean(z,10);

t2=trimmean(z,20);

std(t1)

std(t2)
```

give the standard errors 0.1034 and 0.1004 for  $\bar{X}_{0.1}$  and  $\bar{X}_{0.2}$  respectively.

(e) Matlab commands

c11=prctile(t1,5)c12=prctile(t1,95)c21=prctile(t2,5)c22=prctile(t2,95)

give 90% CIs

```
for \mu_{0.1}: (2\bar{X}_{0.1} - c12; 2\bar{X}_{0.1} - c11) = (14.435; 14.765),
for \mu_{0.2}: (2\bar{X}_{0.2} - c22; 2\bar{X}_{0.2} - c21) = (14.463; 14.784).
```

(f) Matlab commands

```
iqr(x)
median(abs(x-median(x)))
```

Warning: mad(x) in Matlab stands for the mean abs. dev.

(g) Matlab commands (vector z comes from the (d) part) q=prctile(z,75); hist(q) std(q)

give the standard error 0.1332 of the upper quartile.

#### Problem 10.40

Matlab command (x = control and y = seeded data)

qqplot(x,y)

fits the line y = 2.5x claiming 2.5 times more rainfall from seeded clouds. Matlab command

```
qqplot(log(x), log(y))
```

fits the line

 $\log(y) = 2 + 0.8\log(x)$ 

meaning a decreasing slope in the relationship  $y = 7.4x^{0.8}$ .