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## Solutions chapter 13

Warning: in some of the contingency tables the expected counts are rounded. If you then will compute the chi-square test statistic $X^{2}$ from the table, you will often get a somewhat different value.

## Problem 13.1

Test
$H_{0}$ : same genotype frequencies for diabetics and normal using the chi-square test of homogeneity.

|  | Diabetic | Normal | Total |
| :--- | :---: | :---: | :---: |
| $B b$ or $b b$ | $12(7.85)$ | $4(8.15)$ | 16 |
| $B B$ | $39(43.15)$ | $49(44.85)$ | 88 |
| Total | 51 | 53 | 104 |

Observed $X^{2}=5.10, \mathrm{df}=1, \mathrm{P}$-value $P=0.024$. Reject $H_{0}$. Diabetics have genotype $B B$ less often.
The exact Fisher test uses $\operatorname{Hg}\left(104,51, \frac{16}{104}\right)$ as the null distribution of the test statistic $N_{11}=12$
one-sided P-value: 1-hygecdf $(11,104,16,51)=0.0225$,
two-sided P -value $P=0.045$.
Normal approximation of the null distribution

$$
\operatorname{Hg}\left(104,51, \frac{16}{104}\right) \approx \mathrm{N}(7.85,3.41)
$$

Since $Z=\frac{12-7.85}{\sqrt{3.41}}=2.245$, the approximate two-sided $P$-value $P=0.025$.

## Problem 13.3

Incidence of tuberculosis in relation to blood groups.
(a) $H_{0}$ : no association of the disease and the ABO blood group:

|  | O | A | AB | B | Total |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Moderate | $7(10.4)$ | $5(9.8)$ | $3(2.0)$ | $13(6.2)$ | 28 |
| Minimal | $27(30.4)$ | $32(29.7)$ | $8(6.1)$ | $18(18.8)$ | 85 |
| Not present | $55(48.6)$ | $50(47.5)$ | $7(9.8)$ | $24(30.0)$ | 136 |
| Total | 89 | 87 | 18 | 55 | 249 |

Observed $X^{2}=15.37, \mathrm{df}=6, \mathrm{P}$-value $P=0.018$. Reject $H_{0}$.
(b) $H_{0}$ : no association of the disease and the MN blood group:

|  | MM | MN | NN | Total |
| :--- | :---: | :---: | :---: | :---: |
| Moderate | $21(16.7)$ | $6(9.4)$ | $1(1.9)$ | 28 |
| Minimal | $54(51.3)$ | $27(28.9)$ | $5(5.8)$ | 86 |
| Not present | $74(81.1)$ | $51(45.7)$ | $11(9.2)$ | 136 |
| Total | 149 | 84 | 17 | 250 |

Observed $X^{2}=4.73, \mathrm{df}=4$, P -value $P=0.42$. Can not reject $H_{0}$.

## Problem 13.6

Goodness of fit chi-square test for $H_{0}$ : boys proportions $p_{12}=p_{22}=p_{32}=0.513$, same sex ratio for three father's activities. (Here 0.513 is obtained as $105.37 /(105.37+100)=0.513$.)

|  | Girl | Boy | Total |
| :--- | :---: | :---: | :---: |
| Flying fighter | $51(45.15)$ | $38(43.84)$ | 89 |
| Flying transport | $14(15.22)$ | $16(14.78)$ | 30 |
| Not flying | $38(42.62)$ | $46(41.38)$ | 84 |
| Total | 103 | 100 | 203 |

Observed $X^{2}=2.75, \mathrm{df}=3, \mathrm{P}$-value $P=0.43$. Can not reject $H_{0}$.

## Problem 13.8

A randomized double-blind experiment compared the effectiveness of several drugs in ameliorating postoperative nausea. All patients were anesthetized with nitrous oxide and ether. The following table shows the incidence of nausea during the first four hours for each of several drugs and a placebo.

Number of patients Incidence of nausea

Placebo
Chlorpromazine
Dimenhydrinate
Pentobarbital ( 100 mg )
Pentobarbital ( 150 mg )

95
52
52
35
37

We compare the drugs to each other and to placebo using the chi-square test for homogeneity

|  | No nausea | Incidence of nausea | Total |
| :--- | :---: | :---: | :---: |
| Placebo | $70(84)$ | $95(81)$ | 165 |
| Chlorpromazine | $100(78)$ | $52(74)$ | 152 |
| Dimenhydrinate | $33(43)$ | $52(42)$ | 85 |
| Pentobarbital $(100 \mathrm{mg})$ | $32(34)$ | $35(33)$ | 67 |
| Pentobarbital $(150 \mathrm{mg})$ | $48(43)$ | $37(42)$ | 85 |
| Total $(150 \mathrm{mg})$ | 283 | 271 | 554 |

The observed test statistic $X^{2}=35.8$ according to the $\chi_{4}^{2}$-distribution table gives P -value $=3 \cdot 10^{-7}$. Comparing the observed and expected counts we conclude that Placebo and Dimenhydrinate are most effective in ameliorating postoperative nausea.

## Problem 13.17

A study of the relation of blood type to peptic ulcer.
(a) $H_{0}$ : no relation between blood group and disease in London:

|  | Control | Peptic Ulcer | Total |
| :--- | :---: | :---: | :---: |
| Group A | $4219(4103.0)$ | $579(695.0)$ | 4798 |
| Group O | $4578(4694.0)$ | $911(795.0)$ | 5489 |
| Total | 8797 | 1490 | 10287 |

Observed $X^{2}=42.40, \mathrm{df}=1$, P -value $P=0.000$. Reject $H_{0}$. Odds ratio $\hat{\Delta}=1.45$.
(b) $H_{0}$ : no relation between blood group and disease in Manchester:

|  | Control | Peptic Ulcer | Total |
| :--- | :---: | :---: | :---: |
| Group A | $3775(3747.2)$ | $246(273.8)$ | 4021 |
| Group O | $4532(4559.8)$ | $361(333.2)$ | 4893 |
| Total | 8307 | 607 | 8914 |

Observed $X^{2}=5.52, \mathrm{df}=1, \mathrm{P}$-value $P=0.019$. Reject $H_{0}$. Odds ratio $\hat{\Delta}=1.22$.
(c) $H_{0}$ : London Group A and Manchester Group A have the same propensity to Peptic Ulcer:

|  | C and A | PU and A | Total |
| :--- | :---: | :---: | :---: |
| London | $4219(4349.2)$ | $579(448.8)$ | 4798 |
| Manchester | $3775(3644.8)$ | $246(376.2)$ | 4021 |
| Total | 7994 | 825 | 8819 |

Observed $X^{2}=91.3, \mathrm{df}=1$, P-value $P=0.000$. Reject $H_{0}$.
$H_{0}$ : London Group O and Manchester Group O have the same propensity to Peptic Ulcer:

|  | C and O | PU and O | Total |
| :--- | :---: | :---: | :---: |
| London | $4578(4816.5)$ | $911(672.5)$ | 5489 |
| Manchester | $4532(4293.5)$ | $361(599.5)$ | 4893 |
| Total | 9110 | 1272 | 10382 |

Observed $X^{2}=204.5, \mathrm{df}=1$, P-value $P=0.000$. Reject $H_{0}$.

## Problem 13.18

$\mathrm{D}=$ endometrical carcinoma, $\mathrm{X}=$ estrogen taken at least 6 months prior to the diagnosis of cancer.
(a) Matched controls, retrospective case-control study

|  | Controls: estrogen used | Controls: estrogen not used | Total |
| :--- | :---: | :---: | :---: |
| Cases: estrogen used | 39 | 113 | 152 |
| Cases: estrogen not used | 15 | 150 | 165 |
| Total | 54 | 263 | 317 |

Apply McNemar test for

$$
H_{0}: \pi_{1 \cdot}=\pi_{\cdot 1} \quad \text { vs } \quad H_{1}: \pi_{1} \neq \pi_{\cdot 1} .
$$

Observed value of the test statistic

$$
X^{2}=\frac{(113-15)^{2}}{113+15}=75
$$

is highly significant as $\sqrt{75}=8.7$ and the corresponding two-sided P-value obtained from $\mathrm{N}(0,1)$ table is very small.
(b) Possible weak points in a retrospective case-control design

- selection bias: some patients have died prior the study,
- information bias: have to rely on other sources of information.


## Problem 13.19

The effect of anxiety on a person's desire to be alone or in company.

|  | Wait Together | Wait Alone | Total |
| :--- | :---: | :---: | :---: |
| High-Anxiety | 12 | 5 | 17 |
| Low-Anxiety | 4 | 9 | 13 |
| Total | 16 | 14 | 30 |

Test
$H_{0}$ : same frequencies ratio for High-Anxiety and Low-Anxiety groups.
What is a reasonable one-sided alternative?
(a) The exact Fisher test uses $\operatorname{Hg}\left(30,17, \frac{16}{30}\right)$ as the null distribution of the test statistic $N_{11}=12$. It gives
one-sided P-value: 1-hygecdf $(11,30,16,17)=0.036$, two-sided P -value $P=0.071$.
(b) Using normal approximation

$$
\operatorname{Hg}\left(30,17, \frac{16}{30}\right) \approx \mathrm{N}\left(9.1,(1.4)^{2}\right)
$$

and continuity correction, we find the one-sided P -value to be

$$
\mathrm{P}_{H_{0}}\left(N_{11} \geq 12\right)=\mathrm{P}_{H_{0}}\left(N_{11}>11\right) \approx 1-\Phi\left(\frac{11.5-9.1}{1.4}\right)=1-\Phi(1.71)=0.044
$$

(c) Approximate chi-square test yields: observed $X^{2}=4.69, \mathrm{df}=1$, two-sided P-value

$$
2(1-\Phi(\sqrt{4.69}))=2(1-\Phi(2.16))=0.03
$$

## Problem 13.24

Red against blue outfits - does it matter in combat sports? The winners in different sports

|  | Red | Biue | Total |
| :--- | :---: | :---: | :---: |
| Boxing | 148 | 120 | 268 |
| Freestyle wrestling | 27 | 24 | 51 |
| Greco-Roman wrestling | 25 | 23 | 48 |
| Tae Kwon Do | 45 | 35 | 80 |
| Total | 245 | 202 | 447 |

Denote
$\pi_{1}=$ probability that red wins in boxing,
$\pi_{2}=$ probability that red wins in freestyle wrestling,
$\pi_{3}=$ probability that red wins in Greco-Roman wrestling,
$\pi_{4}=$ probability that red wins in Tae Kwon Do.
(a) Assuming

$$
H_{e q}: \pi_{1}=\pi_{2}=\pi_{3}=\pi_{4}=\pi
$$

we test

$$
H_{0}: \pi=\frac{1}{2} \quad \text { vs } \quad H_{1}: \pi \neq \frac{1}{2} .
$$

We use the large sample test for proportion based on the statistic $X=245$ whose null distribution is $\operatorname{Bin}\left(n, \frac{1}{2}\right), n=447$. The two-sided P-value is approximated by

$$
2\left(1-\Phi\left(\frac{245-\frac{447}{2}}{\sqrt{447} \cdot \frac{1}{2}}\right)=2(1-\Phi(2.034)=0.042\right.
$$

At $5 \%$ level we reject the $H_{0}: \pi=\frac{1}{2}$. The MLE is $\hat{\pi}=\frac{245}{447}=0.55$.
(d) Is there evidence that wearing red is more favourable in some of the sports than others? We test

$$
H_{e q}: \pi_{1}=\pi_{2}=\pi_{3}=\pi_{4} \quad \text { vs } \quad H_{\text {ineq }}: \pi_{i} \neq \pi_{j} \quad \text { for some } i \neq j
$$

using the chi-square test of homogeneity. From

|  | Red | Biue | Total |
| :--- | :---: | :---: | :---: |
| Boxing | $148(147)$ | $120(121)$ | 268 |
| Freestyle wrestling | $27(28)$ | $24(23)$ | 51 |
| Greco-Roman wrestling | $25(26)$ | $23(22)$ | 48 |
| Tae Kwon Do | $45(44)$ | $35(36)$ | 80 |
| Total | 245 | 202 | 447 |
| Marginal proportions | 0.55 | 0.45 | 1.00 |

we find that the test statistic $X^{2}=0.3$ is not significant. We can not reject $H_{e q}$, which according to (a) leads to $\hat{\pi}=0.55$.
(b) Now we state the hypotheses of interest directly: consider

$$
H_{0}: \pi_{1}=\pi_{2}=\pi_{3}=\pi_{4}=\frac{1}{2} \quad \text { vs } \quad H_{1}:\left(\pi_{1}, \pi_{2}, \pi_{3}, \pi_{4}\right) \neq\left(\frac{1}{2}, \frac{1}{2}, \frac{1}{2}, \frac{1}{2}\right)
$$

Here we need a new chi-square test, a chi-square test for $k$ proportions with $k=4$. We derive it using the likelihood ratio approach.

Given four observed counts $x_{1}=148, x_{2}=27, x_{3}=25, x_{4}=45$, the likelihood function based on the binomial model has the form

$$
L\left(\pi_{1}, \pi_{2}, \pi_{3}, \pi_{4}\right)=\prod_{i=1}^{4}\binom{n_{i}}{x_{i}} \pi_{i}^{x_{i}}\left(1-\pi_{i}\right)^{n_{i}-x_{i}} .
$$

Using $\hat{\pi}_{i}=\frac{x_{i}}{n_{i}}$, we compute the likelihood ratio as

$$
\Lambda=\frac{L\left(\frac{1}{2}, \frac{1}{2}, \frac{1}{2}, \frac{1}{2}\right)}{L\left(\hat{\pi}_{1}, \hat{\pi}_{2}, \hat{\pi}_{3}, \hat{\pi}_{4}\right)}=\frac{\left(\frac{1}{2}\right)^{n}}{\prod_{i=1}^{4}\left(\frac{x_{i}}{n_{i}}\right)^{x_{i}}\left(\frac{n_{i}-x_{i}}{n_{i}}\right)^{n_{i}-x_{i}}} .
$$

Turning to the logarithms,

$$
\Delta=-\ln \Lambda=\sum_{i=1}^{4} x_{i} \ln \frac{2 x_{i}}{n_{i}}+\left(n_{i}-x_{i}\right) \ln \frac{2\left(n_{i}-x_{i}\right)}{n_{i}}
$$

we take $2 \Delta$ as the test statistic.
Next we show that the null distribution of $2 \Delta$ is approximately $\chi_{4}^{2}$. Under $H_{0}$ we have $\frac{2 x_{i}}{n_{i}} \approx 1$, and using a Taylor expansion we find that

$$
2 \Delta \approx X^{2}=\sum_{i=1}^{4} \frac{\left(x_{i}-\frac{n_{i}}{2}\right)^{2}}{n_{i} / 4},
$$

where $Z_{i}=\frac{X_{i}-\frac{n_{i}}{2}}{\sqrt{n_{i} / 4}}$ are independent and approximately $\mathrm{N}(0,1)$ distributed, provided $X_{i} \sim \operatorname{Bin}\left(n_{i}, \frac{1}{2}\right)$. From

|  | Red | Biue | Total |
| :--- | :---: | :---: | :---: |
| Boxing | $148(134)$ | $120(134)$ | 268 |
| Freestyle wrestling | $27(25.5)$ | $24(25.5)$ | 51 |
| Greco-Roman wrestling | $25(24)$ | $23(24)$ | 48 |
| Tae Kwon Do | $45(40)$ | $35(40)$ | 80 |
| $H_{0}$ proportions | 0.5 | 0.5 | 1.00 |

we find $X_{\mathrm{obs}}^{2}=4.4$. Since $\chi_{4}^{2}(0.1)=7.8$, we do not reject $H_{0}: \pi_{1}=\pi_{2}=\pi_{3}=\pi_{4}=\frac{1}{2}$.

