

## Computer exercise 2

# Distributions in Safety Analysis

In this computer exercise we will encounter some fundamental concepts, firstly, from probability theory: the probability density function, expectation, and variance of a random variable; and, secondly, from statistics: the histogram, the empirical distribution, and parameter estimation. The Gumbel distribution and the Weibull distribution, both often used in safety analysis, will serve as examples. At first we will rely on simulations, but eventually we will investigate real-world data: measurements of wave heights from the Atlantic Ocean. All necessary files are downloadable from the course home page <http://www.math.chalmers.se/Stat/Grundutb/CTH/mve300/1112/files/labfiles.zip>. Please download the labfiles.zip file and uncompress it at the directory you plan to use for the computer exercises.

### 1 Preparatory exercises

**Question 1:** Write down the definitions of expectation and variance of a continuous random variable  $X$ , i.e.  $E(X)$  and  $V(X)$ . Derive the expectation and variance of  $X$  if  $X$  is exponentially distributed.

**Question 2:** Compute the likelihood function  $L(a; \mathbf{x})$  if  $\mathbf{x} = (x_1, \dots, x_n)$  is a sample from an exponential distribution.

### 2 How to generate random numbers

Let  $Y$  be a uniformly distributed random variable (between 0 and 1), and let  $F$  be a distribution function. Then a random variable  $X$  is said to be a random variable distributed according to  $F$  if  $F(X) = Y$ , i.e. if<sup>1</sup>

$$X = F^{-1}(Y).$$

If, for instance,  $F$  is

$$\begin{array}{lll} \text{Weibull}^2: & Y = F(X) = 1 - e^{-\left(\frac{X-b}{a}\right)^c} & \Leftrightarrow X = b + a(-\ln(1-Y))^{1/c} \\ \text{normal:} & Y = F(X) = \Phi\left(\frac{X-m}{\sigma}\right) & \Leftrightarrow X = m + \sigma\Phi^{-1}(Y) \\ \text{Gumbel:} & Y = F(X) = \exp\left(-e^{-\frac{X-b}{a}}\right) & \Leftrightarrow X = b - a\ln(-\ln Y), \end{array}$$

<sup>1</sup>This is a not very precise formulation; please see Section 3.1.2 in the textbook. Note that  $F^{-1}(y)$  is the *inverse* function of  $F$  (at instant  $y$ ), *not* the *reciprocal* value  $1/F(y)$  of  $F(y)$ .

then  $X$  is a Weibull, normally, and Gumbel distributed random variable, respectively. In Matlab uniformly random variables (“random numbers”) are generated by means of the command `rand`. We will use it here to produce 500 Weibull-distributed random-numbers:

```
>> a=2; b=0; c=3.6;
>> x=b+a*(-log(1-rand(500,1))).^(1/c);
>> plot(x, '.'), grid on
```

**Question 3:** Why is there a full stop before the exponent on the second row in the Matlab code here?

or 2000 normally-distributed random numbers:

```
>> m=10; sigma=3;
>> x=m+sigma*norminv(rand(2000,1));
>> plot(x, '.'), grid on
```

Here we *really* encourage you to use the command `randn` instead, i.e

```
>> x=m+sigma*randn(2000,1);
```

Eventually, produce 35 Gumbel-distributed random numbers:

```
>> a=2; b=3.6;
>> x=b-a*log(-log(rand(35,1)));
>> plot(x, '.'), grid on
```

**Question 4:** What do the plots look like? Do you see any regularity? Make any descriptive comments on the type of data you see!

This type of plot may indicate the “average” value and spreading, but in the sequel we will illustrate data graphically in a more convenient way. To generate random numbers in Matlab, one can also make use of the commands `wblrnd` (Weibull), `normrnd` (normal), and `raylrnd` (Rayleigh) from the commercial Statistics Toolbox.

### 3 Probability density function as a limit of histograms

In descriptive statistics the histogram is used as one way to describe the distribution of data. We will now compare the histogram with the probability density function (pdf), often denoted by  $f_X(x)$  if the underlying random variable is  $X$ . In the following numerical example,  $X$  belongs to a Gumbel distribution

$$F_X(x) = \exp\left(-e^{-(x-b)/a}\right)$$

with parameters  $a = 2, 1, b = 1, 7$ .

Generate 1000 observations:

```
>> a=2.1; b=1.7;
>> x=b-a*log(-log(rand(1,1000)));
```

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<sup>2</sup>Here,  $F$  is defined only for  $X > b$ , i.e when  $F(X) > 0$

**Question 5:** Can you briefly explain what we did in the above lines of code?

To generate  $\mathbf{x}$ , we can also make use of the Matlab Statistical Toolbox built-in command `evrnd`:

```
>> x=-evrnd(-b,a,1,1000);
```

The two minuses in the above command are needed because the Matlab uses the minimal value distribution while the Gumbel distribution refers to the maximum value distribution (ambiguity of the word 'extreme'). See Extreme Value Type I Distribution for more details. To avoid confusion, we have created Matlab functions that allow to work with the Gumbel distribution with the same parametrization as in the textbook. Now, we can run

```
>> x=gumbrnd(a,b,1,1000);
```

Next, make a histogram utilising the command `hist`:

```
>> help hist
>> hist(x)
```

Note that the number of observations in each class is presented on the ordinata (y-axis). From theory, we know that a pdf  $f_X(x)$  always has the property  $\int_{-\infty}^{\infty} f_X(x) dx = 1$ . To compare the histogram with the pdf, one has to *scale* the former.

Redraw the histogram using the `bar` function and rescaled height of the bars so the area of the bars adds to one (this explains the presence of the term  $n/(\text{sum}(n)*(x_{\text{out}}(2)-x_{\text{out}}(1)))$  in the code below). In the same figure, draw the theoretical pdf:

```
>> [n,xout]=hist(x);
>> bar(xout,n/(sum(n)*(xout(2)-xout(1))),1)
>> hold on
>> xv=linspace(min(x),max(x),1000);
>> plot(xv,exp(-(xv-b)/a-exp(-(xv-b)/a))/a,'r')
>> hold off
```

**Question 6:** Do you understand what we have done here? Write down the probability density function for  $X$  and identify it in the Matlab code above.

The Matlab routine `evpdf` gives the pdf for a Gumbel-distributed random variable, so it is also possible to write (note that we used self-written `gumbpdf` rather than using Matlab's `evpdf` which has the confusing sign problem mentioned above):

```
>> [n,xout]=hist(x,100); % 100 in 'hist' represents the number of bins for histogram
                        % you may change it to see what happens
>> bar(xout,n/(sum(n)*(xout(2)-xout(1))),1)
>> hold on
>> xv=linspace(min(x),max(x),500);
>> plot(xv,gumbpdf(xv,a,b),'r')
>> hold off
```

**Question 7:** What would happen if you increased the number of generated values in `xv`? (You may check by increasing 1000 to 2000 to 5000 to 10000.)

## 4 Expectation and variance of a random variable

For a random variable  $X$ , the *expectation*, sometimes called the *mean* and denoted  $E(X)$ , gives the value of  $X$  “on average”; if the distribution of  $X$  had been the *mass* distribution of a physical thing, the expectation would have located the centre of gravity of that thing. The *variance*  $V(X)$  (or, rather, the *standard deviation*  $D(X) = \sqrt{V(X)}$ ) of  $X$  can be regarded as a measure of the distribution’s dispersion. For a set of important distributions,  $E(X)$  and  $V(X)$  have been explicitly derived (in terms of the distribution’s parameters) and tabulated, see for example the textbook.

For a given data set  $x_1, \dots, x_n$  (sample), in most cases we do not know the distribution from which the sample is taken, and hence not the mean and variance of that distribution. The sample mean, often denoted  $\bar{x} = (\sum_{i=1}^n x_i)/n$ , and the sample variance, often denoted  $s^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})^2$ , are then the corresponding measures of location and dispersion. If the number  $n$  of observations increases, we may expect that these quantities become closer to  $E(X)$  and  $V(X)$  respectively. Let us examine this in Matlab by means of simulated data, the distribution of which we can control:

Consider the Weibull distribution,

$$F_X(x) = 1 - \exp(-((x - b)/a)^c), \quad x \geq b.$$

The mean and variance are given by

$$\begin{aligned} E(X) &= b + a\Gamma\left(1 + \frac{1}{c}\right), \\ V(X) &= a^2\Gamma\left(1 + \frac{2}{c}\right) - a^2\left(\Gamma\left(1 + \frac{1}{c}\right)\right)^2, \end{aligned}$$

where

$$\Gamma(p) = \int_0^{\infty} x^{p-1}e^{-x} dx. \quad (1)$$

is the gamma function. Choose for example  $a = 1.5$ ,  $b = 0$ , and  $c = 2$ . To calculate expectation and variance, one needs the gamma function in (1) which is implemented in Matlab as `gamma`; hence

```
>> a=1.5; b=0; c=2;
>> EX=b+a*gamma(1+1/c)
>> VX=a^2*gamma(1+2/c)-a^2*(gamma(1+1/c))^2;
>> DX=sqrt(VX)
```

Now, simulate a sample of 50 observations and find the sample mean and standard deviation by the commands `mean` and `std` respectively:

```
>> x=b+a*(-log(1-rand(1,50))).^(1/c);
>> mean(x), std(x)
```

(Again, can you understand this simulation?) Since  $b = 0$ , alternatively you can also use the Matlab built-in routine `wblrnd`

```
>> x=wblrnd(a,c,1,50);
>> mean(x), std(x)
```

**Question 8:** Compare the values estimated from the samples with the theoretical values  $EX$ ,  $DX$  that you have also obtained above. Write down the values for  $E(X)$ ,  $\bar{x}$ ,  $D(X)$ ,  $d(x)$ . Are the theoretical and empirical values consistent with each other? Simulate larger samples of, say, 200, 1000, and 5000 observations respectively. What happens when the number of observations increases?

## 5 Estimation of parameters

Assume that we have a sample  $x_1, \dots, x_n$  from (for example) a Gumbel distribution, i.e. the distribution function is

$$F(x) = \exp\left(-e^{-(x-b)/a}\right).$$

However, the parameters  $a$  and  $b$  are not known. Then, one can use the *maximum-likelihood method* (ML method) to estimate the parameters from the sample.

**Question 9:** Write down the likelihood function  $L(a, b; \mathbf{x})$  for the example above.

In the Statistical Toolbox the ML method has been implemented in `evfit`, `wblfit`, and `raylfit` for the purpose of estimating the parameters in a Gumbel, Weibull, and Rayleigh distribution respectively.

First, simulate a sample of, say, 50 observations from a Gumbel distribution, then check if the ML method implemented in `gumbfit` (which is based on `evfit`) returns good estimates:

```
>> a=2; b=3.5;
>> x=b-a*log(-log(rand(1,50))); % Alternative 1
>> x=gumbrnd(a,b,1,50); % Alternative 2
>> phat= gumbfit(x);
>> hata=phat(1)
>> hatb=phat(2)
```

**Question 10:** The parameter estimates are given in the vector `phat`. With the elements `phat(1)` and `phat(2)` corresponding to  $\hat{a}$  and  $\hat{b}$ , respectively. Compare the estimates  $\hat{a}$  and  $\hat{b}$  with the true values  $a$  and  $b$ ! Write them down.

### Properties of point estimates

The variances and covariances of the point estimates are always of interest. For the point estimates  $\hat{a}$  and  $\hat{b}$  of  $a$  and  $b$  in a Gumbel distribution above, the asymptotic variances and covariance (when the number of observations “large”) are given by<sup>3</sup>

$$\begin{aligned} V(\hat{a}) &\approx \frac{6}{\pi^2} \cdot \frac{a^2}{n} \approx 0,607\,93 \cdot \frac{a^2}{n} \\ V(\hat{b}) &\approx \left(1 + \frac{6(1-\gamma)^2}{\pi^2}\right) \cdot \frac{a^2}{n} \approx 1,108\,67 \cdot \frac{a^2}{n} \\ C(\hat{a}, \hat{b}) &\approx \frac{6(1-\gamma)}{\pi^2} \cdot \frac{a^2}{n} \approx 0,257\,02 \cdot \frac{a^2}{n} \end{aligned}$$

**Question 11:** Evaluate estimates of  $V(\hat{a})$ ,  $V(\hat{b})$ , and  $C(\hat{a}, \hat{b})$ , using  $\hat{a}$  instead of  $a$ , report the obtained values.

<sup>3</sup> It is not at all trivial to show this. Here,  $\gamma \stackrel{\text{def}}{=} \lim_{k \rightarrow \infty} \left(\sum_{i=1}^k (1/i) - \ln k\right) = 0,577\,215\,665\dots$  is Euler’s constant (or the Euler-Mascheroni constant to distinguish from  $e$  which is also frequently referred to as Euler’s constant); it is not known whether  $\gamma$  is irrational or not (the seventh Hilbert problem)!

## Estimation of quantiles

By means of  $\hat{a}$  and  $\hat{b}$ , estimate the upper 1 % quantile, defined as the number  $x_{0,01}$  which satisfies

$$P(X > x_{0,01}) = 0,01 \Leftrightarrow 1 - F(x_{0,01}) = 0,01.$$

Thus, the equation

$$1 - \exp\left(-e^{-(x_{0,01}-b)/a}\right) = 0,01$$

must be solved with respect to  $x_{0,01}$ ; we obtain

$$x_{0,01} = b - a \ln(-\ln(1 - 0,01)).$$

A reasonable estimate  $\widehat{x_{0,01}}$  of  $x_{0,01}$  would then be

$$\widehat{x_{0,01}} = \hat{b} - \hat{a} \ln(-\ln(1 - 0,01)). \quad (2)$$

So it can be obtained from

```
>> xhat=phat(2)-phat(1)*log(-log(1-0.01))
```

**Question 12:** Get a numerical result (from the above) for the estimate of the quantile.

Since  $\hat{a}$  and  $\hat{b}$  both are random variables, so is  $\widehat{x_{0,01}}$  according to Equation (2). Then  $\widehat{x_{0,01}}$  possesses an expectation  $E(\widehat{x_{0,01}})$  and a standard deviation  $D(\widehat{x_{0,01}})$ . The standard deviation indicates the dispersion of the estimate  $\widehat{x_{0,01}}$ , and it is therefore important to get an idea of the value of  $D(\widehat{x_{0,01}})$ . In most cases it is impossible to find an exact value, and consequently an approximation has to do. Such an approximation is called a *standard error*. By letting  $Z_1 = \hat{b}$ ,  $Z_2 = \hat{a}$ ,  $c_1 = 1$ , and  $c_2 = -\ln(-\ln(1 - 0,01))$ , we can make use of the formula

$$V(c_1 Z_1 + c_2 Z_2) = c_1^2 V(Z_1) + c_2^2 V(Z_2) + 2c_1 c_2 C(Z_1, Z_2)$$

to obtain a standard error for  $\widehat{x_{0,01}}$ :

$$\begin{aligned} D(\widehat{x_{0,01}}) &= \sqrt{V(\hat{b} - \hat{a} \ln(-\ln(1 - 0,01)))} = \\ &= \sqrt{V(\hat{b}) + (-\ln(-\ln(0,99)))^2 \cdot V(\hat{a}) + 2 \cdot (-\ln(-\ln(0,99))) \cdot C(\hat{b}, \hat{a})} \end{aligned}$$

**Question 13:** Use approximations of  $V(\hat{a})$ ,  $V(\hat{b})$ , and  $C(\hat{b}, \hat{a})$  you got in Question 11 to obtain the estimate of the standard deviation (aka standard error). Write down its numerical value. Is it large comparing to the estimate value of quantile?

Of course, just 50 values to estimate  $x_{0,01}$  might be too small a number leading to a quite large standard error.

## 6 Probability plots

Assume that we have a set of observations  $x_1, x_2, \dots, x_n$ . Before we estimate any parameters, we must convince ourselves that the observations originate from the right *family* of distributions, e.g. normal, Gumbel, or Weibull. One way to get a rough idea of which family of distributions may be suitable, is to display the observations in a *probability plot*<sup>4</sup>: If you suspect that the data originate from, for instance, a normal distribution, then you should make a *normal probability plot*; if you instead suspect a Gumbel distribution, then make a *Gumbel probability plot*. If, in the plot, the observations seem to line up well along a straight line, it indicates that the chosen distribution for the probability plot indeed might serve as a good model for the observations. Statistics Toolbox provides `normplot` (for normal distribution), `wblplot` (for Weibull distribution); but unfortunately there is no probability plot for Gumbel distribution, so we have created one and named it `gumbplot` (available in `labfiles`). Acquaint yourself with the above-mentioned commands, for example

```
>> dat1=randn(2000,1); % Attention: Normal distribution!
>> normplot(dat1)
>> wblplot(dat1)
>> dat2=rand(3000,1); % Attention: Uniform distribution!
>> normplot(dat2)
>> gumbplot(dat2)
>> dat3=wblrnd(2,2.3,1,3000); % Attention: Weibull distribution!
>> wblplot(dat3)
>> gumbplot(dat3) % Attention: Gumbel distribution!
>> dat4=gumbrnd(1,2,1,3500); % Available in labfiles
>> gumbplot(dat4)
```

Experiment more with the number of observations; change also distributions!

**Question 14:** What happens when you plot the data in the “wrong” distribution plot?

### Measurements of significant wave heights in the Atlantic Ocean

In the field of oceanography and marine technology, statistical extreme-value theory has been used to a great extent. In design of offshore structures knowledge about “extreme” conditions is important.

In the numerical examples above, we used artificial data, simulated from a distribution which we could control. We will now consider *real* measurements from the Atlantic Ocean. The data set contains so-called significant wave heights (in meters), that is, the average of the highest one-third of the waves.

Now, load the data set `atlantic.dat` and read about the measurements; then find the size of data, and plot it:

```
>> atl=load('atlantic.dat');
>> help atlantic
>> size(atl)
>> plot(atl, '.')
```

---

<sup>4</sup>Before the computer age, the observations were plotted manually into diagram-forms printed on sheets of paper; therefore we now and then will use the expression “to plot data in a certain probability paper” even if we are referring to computer-displayed diagrams.

One knows that, roughly speaking, the registered so-called significant wave-heights behave, statistically, as if they were maximum wave-heights; therefore one can suspect them to originate from a Gumbel distribution, for instance. Below we will make different probability plots.

```
>> normplot(at1)
>> normplot(log(at1))
>> gumbplot(at1)
>> wblplot(at1)
```

<p><b>Question 15:</b> Which distribution might be a satisfactory choice? Estimate parameters as in Section 5 for the distribution of your choice (<code>gumbfit</code>, <code>wblfit</code>, <code>normfit</code>).</p>
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