

# Lecture 8. Conditional Distributions - introduction to Bayesian Inference<sup>1</sup>

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<sup>1</sup>Bayesian statistics is a general methodology to analyse and draw conclusions from data.

## The conditional cdf $P(X \leq x|Y = y)$ . and pdf

Suppose that we observed the value of  $Y$ , e.g. we know that  $Y = y$ , but  $X$  is not observed yet. An important question is if the uncertainty about  $X$  is affected by our knowledge that  $Y = y$ , i.e. if

$$F(x|y) = P(X \leq x|Y = y)$$

depends on  $y^2$ .

For continuous r.v.  $X, Y$  it is not obvious how to define conditional probabilities given that " $Y = y$ ", since  $P(Y = y) = 0$  for all  $y$ . It is done using the conditional probability density

$$f(x|y) = \frac{f(x, y)}{f(y)}, \quad F(x|y) = \int_{-\infty}^x f(\tilde{x}|y) d\tilde{x}$$

is the conditional distribution.

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<sup>2</sup>If  $X$  and  $Y$  are independent then obviously  $F(x|y) = F_X(x)$  and  $Y$  gives us no knowledge about  $X$ .

## Law of Total Probability - continuous case

If  $X$  and  $Y$  have joint density  $f(x, y)$  and  $B$  is a statement about  $X$ , then

$$P(B) = \int_{-\infty}^{+\infty} P(B|Y = y)f_Y(y) dy. \quad P(B | Y = y) = \int_B f(x|y) dx,$$

### Example 6

In remote located scientific station there is a supply of food for  $Y \in \text{Exp}(a_Y)$  days. Waiting time for the new delivery is  $X \in \text{Exp}(a_X)$ . Compute probability that food will not finish, i.e.  $P(X < Y)$  assuming that  $X, Y$  are independent.

## Bayes Formula

In many examples the new piece of information is formulated in form of a statement that is true. For example let  $Y$  be strength of a wire and let  $C$  = "the wire passed preloading test of 1000kg", i.e.  $C = "Y > 1000"$  is true. If the likelihood  $L(y) = P(C|Y = y)$  is known then the density  $f(y|C)$  is computed using Bayes formula

$$f_Y^{pos}(y) = f(y|C) = cP(C|Y = y)f(y), \quad c = 1/P(C).$$

Law of total probability gives  $P(C) = \int_0^{\infty} P(C|Y = y)f(y) dy$ .

## Typical problem in safety of existing structure:

Suppose a wire has strength  $Y$  and that loads during years  $i$ ,  $X_i$ , are independent. We already can compute

$$P(\text{"wire survives first years load"}) = P(X_1 < Y) = \int_0^{\infty} F_{X_1}(y) f_Y(y) dy.$$

Suppose  $B = \text{"wire survives first years load"}$  is true. What is probability

$$P(\text{"wire survives second year load"}) = \int_0^{\infty} F_{X_2}(y) f_Y^{post}(y) dy.$$

We need the posterior strength  $f_Y^{post}(y)$ , i.e. derivative of  $F_Y(y|B)$ .

### Example 6

Do computations for unrealistic case that  $X_i$  and  $Y$  are exponentially distributed.

## Bayesian methods in risk evaluation - example:

In the following we shall be mostly interested in studying uncertainties in estimation of probabilities in the following setup. The "initiation" events  $A$  are defined and their concurrences are modeled by Poisson point process with intensity  $\lambda_A$ . In order for  $A$  to develop to an accident or catastrophe, some other unfortunate circumstances, described by event  $B$ , have to take place ( $B$  is called a "scenario"). For example, if  $A$  is "fire ignition"  $B$  could be "failure of sprinkler system".

Sometimes one needs multi-scenario event  $B$ , i.e.  $B = B_1 \cup B_2$  where  $B_1, B_2$ , are excluding. The important parameters are  $\lambda_A$ ,  $p_1 = P(B_1)$  and  $p_2 = P(B_2)$ .

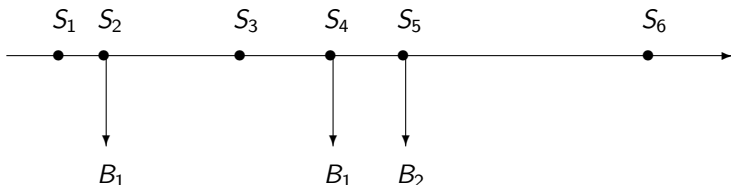


Figure : Events  $A$  at times  $S_i$  with related scenarios  $B_i$ .

$$P_t = P(\text{no accident in period } t) = 1 - e^{-\lambda_A P(B) t} \approx \lambda_A P(B) t,$$

if probability  $P_t$  is small. Hence Two problems of interest in risk analysis:

- ▶ The first one will deal with the estimation of a probability  $p_B = P(B)$ , say, of some event  $B$ , for example the probability of failure of some system.
- ▶ The second one is estimation of the probability that at least once an event  $A$  occurs in a time period of length  $t$ . The problem reduces itself to estimation of the intensity  $\lambda_A$  of  $A$ .

In general parameters  $p_B$  and  $\lambda_A$  are attributes of some physical system, e.g. if  $B =$  "A water sample passes tests" then  $p_B = P(B)$  is a measure of efficiency of a waste-water cleaning process. The intensity  $\lambda_A$  of accidents may characterize a particular road crossing. The parameters  $p_B$  and  $\lambda_A$  are unknown.

## Odds for parameters

Let  $\theta$  denote the unknown value of  $p_B$ ,  $\lambda_A$  or any other quantity.

Introduce odds  $q_\theta$ , which for any pair  $\theta_1, \theta_2$  represents our belief which of  $\theta_1$  or  $\theta_2$  is more likely to be the unknown value of  $\theta$ , *i.e.*  $q_{\theta_1} : q_{\theta_2}$  are odds for the alternatives  $A_1 = " \theta = \theta_1 "$  against  $A_2 = " \theta = \theta_2 "$ .

Since there are here uncountable number of alternatives, we require that  $q_\theta$  integrates to one and hence  $f(\theta) = q_\theta$  is a probability density function representing our belief about the value of  $\theta$ .



## Prior odds - posterior ods

Again, let  $\theta$  be the unknown parameter, for example  $\theta = p_B$ ,  $\theta = \lambda_A$ , while  $\Theta$  denotes any of the variables  $P$  or  $\Lambda$ . Since  $\theta$  is unknown, it is seen as a value taken by a random variable  $\Theta$  with pdf  $f(\theta)$ .

If  $f(\theta)$  is chosen on basis of experience without including observations of outcomes of an experiment then the density  $f(\theta)$  is called a *prior density* and denoted by  $f^{\text{prior}}(\theta)$ .

However, as time passes, our knowledge may change, especially if we observe some outcomes of the experiment which can influence our opinions about the values of parameter  $\theta$  reflecting in the new density  $f(\theta)$ . The modified density  $f(\theta)$  will be called the *posterior density* and denoted by  $f^{\text{post}}(\theta)$ .

The method to update  $f(\theta)$  is

$$f^{\text{post}}(\theta) = cL(\theta) f^{\text{prior}}(\theta)$$

How to find likelihood function  $L(\theta)$  will be discussed later on.

## Predictive probability

Suppose  $f(p)$  has been selected and denote by  $P$  a random variable having pdf  $f(p)$ . A plot of  $f(p)$  is an illustrative measure of how likely the different values of  $p_B$  are.

If only one value of the probability is needed, the Bayesian methodology proposes to use the so-called **predictive probability** which is simply the mean of  $P$ :

$$p^{\text{pred}}(B) = E[P] = \int pf(p) dp.$$

The predictive probability measures the likelihood that  $B$  occurs in future. It combines two sources of uncertainty: the unpredictability whether  $B$  will be true in a future accident and the uncertainty in the value of probability  $p_B$ .

*Example 6.1*

## Predictive probability

As before, if only one single value of the probability is needed, the Bayesian approach proposes to use the predictive probability

$$\begin{aligned} P_t^{\text{pred}}(A) &= E[P] = \int (1 - \exp(-\lambda t)) f_\lambda(\lambda) d\lambda \\ &\approx \int t\lambda f_\lambda(\lambda) d\lambda = tE[\Lambda].^3 \end{aligned}$$

This is a measure of the risk that  $A$  occurs, combining two sources of uncertainty: the variability of the Poisson process of  $A$  and the uncertainty in the intensity of accidents  $\lambda_A$ .

In some situations  $A$  is an initiation event (accident at a crossing) while  $B$  is scenario, e.g.  $B = \text{"Victim needs hospitalisation"}$ . The intensity of  $A \cap B$  is  $\lambda_A P(B)$ . Uncertainty of  $\lambda_A P(B)$  is modeled by  $\Lambda \cdot P$ . The predictive probability of no serious accident is

$$\begin{aligned} P_t^{\text{pred}}(A \cap B) &= \int (1 - \exp(-p\lambda t)) f_\lambda(\lambda) f_P(p) d\lambda dp \\ &\approx \int t p \lambda f_\lambda(\lambda) d\lambda dp = tE[\Lambda]E[P]. \end{aligned}$$