TMS 165/MSA350 Stochastic Calculus Part I Fall 2013 Exercise session 5

Througout this exercise session $B = \{B(t)\}_{t \geq 0}$ denotes Brownian motion.

Exercise 1. Show that

$$X(t) = e^{-\alpha t} \left(\frac{\sigma}{\sqrt{2\alpha}} \left(B(e^{2\alpha t}) - B(1) \right) + x_0 \right) \quad \text{for } t \ge 0$$

is an Ornstein-Uhlenbeck process in the sense that it got the same distributional properties (finite dimensional distributions) as the solution

$$\{X(t)\}_{t\geq 0} = \left\{ e^{-\alpha t} \left(x_0 + \sigma \int_0^t e^{\alpha r} dB(r) \right) \right\}_{t\geq 0}$$

to the Langevin SDE

$$dX(t) = -\alpha X(t) dt + \sigma dB(t)$$
 for $t > 0$, $X(0) = x_0$,

where $\alpha, \sigma > 0$ and $x_0 \in \mathbb{R}$ are constants.

Solution. As both the above X processes are Gaussian they have the same finite dimensional distributions if their mean and covariance functions agree. Here we clearly have $\mathbf{E}\{X(t)\} = \mathrm{e}^{-\alpha t}x_0$ for $t \geq 0$ for both the X processes. Further, we have

$$\mathbf{Cov}\{X(s), X(t)\} = \frac{\sigma^2}{2\alpha} e^{-\alpha(s+t)} \mathbf{Cov}\{B(e^{2\alpha s}) - B(1), B(e^{2\alpha t}) - B(1)\}$$

$$= \frac{\sigma^2}{2\alpha} e^{-\alpha(s+t)} \left(e^{2\alpha \min\{s,t\}} - 1 - 1 + 1\right)$$

$$= \frac{\sigma^2}{2\alpha} \left(e^{-\alpha|s-t|} - e^{-\alpha(s+t)}\right) \quad \text{for } s, t \ge 0$$

for the first X process, while Theorem 4.11 in Klebaner's book shows that

$$\mathbf{Cov}\{X(s),X(t)\} = \sigma^2 e^{-\alpha(s+t)} \int_0^{\min\{s,t\}} \mathrm{e}^{2\alpha r} \, dr = \frac{\sigma^2}{2\alpha} \left(\mathrm{e}^{-\alpha|s-t|} - \mathrm{e}^{-\alpha(s+t)} \right)$$

for $s, t \ge 0$ for the second X process.

Exercise 2. Use the expression for an Ornstein Uhlenbeck process expressed in terms of B from Exercise 1 to find the transition density function for the solution to the Langevin SDE (the Ornstein Uhlenbeck process).

Solution. We have

$$\begin{split} X(t+s) &= \mathrm{e}^{-\alpha(t+s)} \bigg(\frac{\sigma}{\sqrt{2\alpha}} \left(B(\mathrm{e}^{2\alpha(t+s)}) - B(1) \right) + x_0 \bigg) \\ &= \mathrm{e}^{-\alpha(t+s)} x_0 + \frac{\sigma}{\sqrt{2\alpha}} \, \mathrm{e}^{-\alpha(t+s)} \Big(\left(B(\mathrm{e}^{2\alpha(t+s)}) - B(\mathrm{e}^{2\alpha s}) \right) + \left(B(\mathrm{e}^{2\alpha s}) - B(1) \right) \Big) \\ &= \frac{\sigma}{\sqrt{2\alpha}} \, \mathrm{e}^{-\alpha(t+s)} \big(B(\mathrm{e}^{2\alpha(t+s)}) - B(\mathrm{e}^{2\alpha s}) \big) + \mathrm{e}^{-\alpha t} X(s), \end{split}$$

where

$$\frac{\sigma}{\sqrt{2\alpha}} e^{-\alpha(t+s)} \left(B(e^{2\alpha(t+s)}) - B(e^{2\alpha s}) \right)$$

is an N(0, $(\sigma^2/(2\alpha))$ (1-e^{-2 αt}))-distributed random variable independent of $\{X(r)\}_{r\leq s}$. It follows that (X(t+s)|X(s)=x) is N(e^{- αt}x, $(\sigma^2/(2\alpha))$ (1-e^{-2 αt}))-distributed, so that

$$p(y,t+s,x,s) = \frac{d}{dy} P(y,t+s,x,s) = \frac{\sqrt{\alpha}}{\sqrt{\pi (1 - \mathrm{e}^{-2\alpha t})} \sigma} \exp \left\{ -\frac{\alpha (y - x \, \mathrm{e}^{-\alpha t})^2}{\sigma^2 (1 - \mathrm{e}^{-2\alpha t})} \right\}$$

for $t+s > s \ge 0$ and $x, y \in \mathbb{R}$.

Exercise 3. Solve the Stratanovich SDE

$$dX(t) = -\alpha dt + \sigma X(t) \partial B(t)$$
 for $t > 0$, $X(0) = x_0$,

where $\alpha, \sigma > 0$ and $x_0 \in \mathbb{R}$ are constants.

Solution. By Theorem 5.20 in Klebaner's book the above SDE is equivalent to the Itô SDE

$$dX(t) = \left(\frac{1}{2}\sigma^2 X(t) - \alpha\right)dt + \sigma X(t) dB(t) \quad \text{for } t > 0, \quad X(0) = x_0.$$

This in turn is a rather simple form of the linear SDE treated in Section 5.3 in Klebaner's book, with a solution given by

$$X(t) = U(t) \left(x_0 - \alpha \int_0^t \frac{ds}{U(s)} \right)$$
 where $U(t) = e^{\sigma B(t)}$,

which is to say that

$$X(t) = x_0 e^{\sigma B(t)} - \alpha e^{\sigma B(t)} \int_0^t e^{-\sigma B(s)} ds \quad \text{for } t \ge 0.$$

Exercise 4. The CKLS (Chan-Koralyi-Longstaff-Sanders) SDE is given by

$$dX(t) = (\alpha + \beta X(t)) dt + \sigma X(t)^{\gamma} dB(t)$$
 for $t > 0$, $X(0) = x_0$,

where $\alpha, \sigma, \gamma, x_0 > 0$ and $\beta \in \mathbb{R}$ are constants. This SDE is used in contemporary mathematical finance research as a model for, e.g., interest rates and/or deseasonalized eletricity prices, and is famous for being very hard to do inference for and very hard to simulate when $\gamma > 1$. Determine the stationary distribution for this SDE when it exists.

Solution. First note that the fact that $\alpha, x_0 > 0$ ensures that the solution is strictly positive when it exists. From Equation 6.69 in Klebaner's book we further see that the stationary probability density function is given by

$$\pi(x) = \frac{1}{C x^{2\gamma}} \exp \left\{ \int_1^x \frac{2(\alpha + \beta y)}{\sigma^2 y^{2\gamma}} dy \right\} \quad \text{for } x > 0,$$

whenever this function can be normalized to become a density, that is, whenever

$$C = \int_0^\infty \frac{1}{x^{2\gamma}} \exp \left\{ \int_1^x \frac{2(\alpha + \beta y)}{\sigma^2 y^{2\gamma}} \, dy \right\} dx < \infty.$$

The issue whether C is finite or not in turn clearly boils down to check the integrability properties of the function

$$f(x) = \frac{1}{x^{2\gamma}} \exp\left\{ \int_{1}^{x} \frac{2(\alpha + \beta y)}{\sigma^{2} y^{2\gamma}} dy \right\}$$

as $x \downarrow 0$ and as $x \uparrow \infty$. Now, as $x \downarrow 0$ we see that

and as
$$x\uparrow\infty$$
. Now, as $x\downarrow 0$ we see that
$$f(x)\sim \begin{cases} C_1\,x^{-2\gamma} & \text{for } \gamma\in(0,1/2),\\ C_2\,x^{2\alpha/\sigma^2-1} & \text{for } \gamma=1/2,\\ C_3\,x^{-2\gamma}\exp\{-(2\alpha/(\sigma^2(2\gamma-1)))\,x^{-(2\gamma-1)}\} & \text{for } \gamma>1/2, \end{cases}$$

where $C_1, C_2, C_3 > 0$ are constants. This is to say that we always have the integrability required as $x \downarrow 0$. When $x \uparrow \infty$ we further see that

$$f(x) \sim \begin{cases} C_4 x^{-2\gamma} & \text{for } \gamma > 1, \\ C_5 x^{-2+2\beta/\sigma^2} & \text{for } \gamma = 1, \end{cases}$$

$$C_6 x^{-2\gamma} \exp\{(\beta/(\sigma^2(1-\gamma)))x^{2-2\gamma}\} & \text{for } \gamma \in (1/2, 1), \end{cases}$$

$$C_7 x^{2\alpha/\sigma^2 - 1} \exp\{(2\beta/\sigma^2)x\} & \text{for } \gamma = 1/2, \end{cases}$$

$$C_8 x^{-2\gamma} \exp\{(\beta/(\sigma^2(1-\gamma)))x^{2-2\gamma} + (2\alpha/(\sigma^2(1-2\gamma)))x^{1-2\gamma}\} & \text{for } \gamma \in (0, 1/2), \end{cases}$$

where $C_4, \ldots, C_8 > 0$ are constants. This is to say that we have the integrability required when

$$\gamma > 1 \quad \text{and} \quad \gamma = 1, \ 2\beta < \sigma^2 \quad \text{and} \quad \gamma \in (1/2, 1), \ \beta \leq 0 \quad \text{and} \quad \gamma \in (0, 1/2], \ \beta < 0.$$

Exercise 5. Exercise 6.10 in Klebaner's book.

Solution. See the solution at the end of Klebaner's book.

Exercise 6. Let X be a standard normal distributed random variable. Show how X can be made to have any given probability density function $f: \mathbb{R} \to [0, \infty)$ by means of a change of probability measure. Also, if X has probability density function $f: \mathbb{R} \to [0, \infty)$, is it possible to make X have standard normal distributed by a change of probability measure?

Solution. Clearly X has probability density function f under the probability measure

$$\mathbf{Q}(A) = \int_{A} f(X) \sqrt{2\pi} \, \mathrm{e}^{X^{2}/2} \, d\mathbf{P} \quad \text{for } A \in \mathcal{F},$$

as this gives

$$\mathbf{Q}\{X \in B\} = \mathbf{E}_{\mathbf{Q}}\{I_{\{X \in B\}}\}$$

$$= \mathbf{E}_{\mathbf{P}}\{I_{\{X \in B\}} f(X) \sqrt{2\pi} e^{X^{2}/2}\}$$

$$= \int_{\mathbb{R}} I_{B}(x) f(x) \sqrt{2\pi} e^{x^{2}/2} \frac{1}{\sqrt{2\pi}} e^{-x^{2}/2} dx$$

$$= \int_{B} f(x) dx \quad \text{for } B \subseteq \mathbb{R}.$$

If X instead has a strictly positive probability density function $f: \mathbb{R} \to (0, \infty)$ from the beginning, then X is standard normal distributed under the probability measure

$$\mathbf{Q}(A) = \int_A \frac{1}{\sqrt{2\pi}} e^{-X^2/2} \frac{1}{f(X)} d\mathbf{P} \quad \text{for } A \in \mathcal{F},$$

as this gives

$$\begin{aligned} \mathbf{Q}\{X \in B\} &= \mathbf{E}_{\mathbf{Q}} \Big\{ I_{\{X \in B\}} \Big\} \\ &= \mathbf{E}_{\mathbf{P}} \Big\{ I_{\{X \in B\}} \, \frac{1}{\sqrt{2\pi}} \, \mathrm{e}^{-X^2/2} \, \frac{1}{f(X)} \, \Big\} \\ &= \int_{\mathbb{R}} I_B(x) \, \frac{1}{\sqrt{2\pi}} \, \mathrm{e}^{-x^2/2} \, \frac{1}{f(x)} \, f(x) \, dx \\ &= \int_{B} \frac{1}{\sqrt{2\pi}} \, \mathrm{e}^{-x^2/2} \, dx \quad \text{for } B \subseteq \mathbb{R}. \end{aligned}$$

If f is not strictly positive, then it is not possible to make X standard normal distributed by means of this approach, as we then have

$$\begin{aligned} \mathbf{Q}\{\Omega\} &= \mathbf{E}_{\mathbf{Q}}\{1\} \\ &= \mathbf{E}_{\mathbf{P}} \left\{ \frac{1}{\sqrt{2\pi}} e^{-X^2/2} \frac{1}{f(X)} \right\} \\ &= \int_{\{x \in \mathbb{R}: f(x) > 0\}} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} \frac{1}{f(x)} f(x) dx \\ &= \int_{\{x \in \mathbb{R}: f(x) > 0\}} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx \\ &< 1, \end{aligned}$$

so that $\mathbf Q$ is no longer a probability measure.