

Suggested solution for trialexam in
 MSA100/MVE185 Computer Intensive Statistical Methods, October 2008

1. The three observations 2,3,4 have the same likelihood as one normally distributed observation with value 3, and variance $0.5^2/3 = 1/12$. Thus we have a prior distribution with expectation 4 and precision $1/1^2 = 1$, and a likelihood with observed value 3 and precision 12. The posterior becomes a normal distribution with expectation

$$\frac{4 \cdot 1 + 3 \cdot 12}{1 + 12} = \frac{40}{13} = 3.077$$

and standard deviation

$$\sqrt{\frac{1}{1 + 12}} = \sqrt{\frac{1}{13}} = 0.277$$

2. If we are studying the probability distribution for θ , if we know that its density $\pi(\theta)$ is proportional to the function $g(\theta)$, and if $g(\theta)$ is approximated by the density $p(\theta)$ from which we know how to simulate, then importance sampling can be used to find an approximation for the expected value of some function h of the parameter θ . This is done by simulating $\theta_1, \theta_2, \dots, \theta_n$ from the density p , computing the function $h(\theta_i)$ of each of these, and taking the weighted average of these values, with weights $g(\theta_i)/p(\theta_i)$. In contrast, sampling importance resampling is used to find an approximate sample from the distribution π . The θ 's and weights are computed as above, but now the weights are used to find probabilities used for a resampling from the θ 's, in order to obtain an approximate sample.
3. We get

$$\begin{aligned} \pi(\lambda | \theta) &= 0.8 \frac{1}{4} I(|\lambda_1 - \theta_1| \leq 1) I(|\lambda_2 - \theta_2| \leq 1) \\ &\quad + 0.2 \frac{1}{|2\pi\Sigma|^{1/2}} \exp\left(-\frac{1}{2} \lambda' \Sigma^{-1} \lambda\right) \\ &= 0.2 I(|\lambda_1 - \theta_1| \leq 1) I(|\lambda_2 - \theta_2| \leq 1) \\ &\quad + 0.2 \frac{1}{2\pi 10} \exp\left(-\frac{1}{20} (\lambda_1^2 + \lambda_2^2)\right) \end{aligned}$$

so for the acceptance probability, we get

$$\begin{aligned} r &= \min\left(1, \frac{g(\lambda)/\pi(\lambda | \theta)}{g(\theta)/\pi(\theta | \lambda)}\right) \\ &= \min\left(1, \frac{g(\lambda) \left(20\pi I(|\lambda_1 - \theta_1| \leq 1) I(|\lambda_2 - \theta_2| \leq 1) + \exp(-\frac{1}{20}(\theta_1^2 + \theta_2^2))\right)}{g(\theta) \left(20\pi I(|\lambda_1 - \theta_1| \leq 1) I(|\lambda_2 - \theta_2| \leq 1) + \exp(-\frac{1}{20}(\lambda_1^2 + \lambda_2^2))\right)}\right) \end{aligned}$$

4. We use the notation in the lecture notes. With the Sidák adjustment, $f_i(T(\theta))$ depends only on $T_i(\theta)$, so if all the test statistics are independent, then so are $f_i(T(\theta))$ for all i , and thus so are all

$$v_i = H_i(\theta)(1 - f_i(T(\theta)))$$

for all i . Also, $E(f_i(T(\theta))) = \Pr(1 - (1 - T_i(\theta))^N > \alpha) = (1 - \alpha)^{1/N}$. So we get

$$\begin{aligned}
FWER &= \Pr(V > 0) \\
&= 1 - E(I(V = 0)) \\
&= 1 - E\left(\prod_{i=1}^N (1 - v_i)\right) \\
&= 1 - \prod_{i=1}^N (1 - E(v_i)) \\
&= 1 - \prod_{i=1}^N (1 - H_i(\theta)(1 - (1 - \alpha)^{1/N})) \\
&\leq 1 - \prod_{i=1}^N (1 - (1 - (1 - \alpha)^{1/N})) \\
&= \alpha
\end{aligned}$$

5. (a) For the prior, we have $\pi(p) = 1$, so that

$$\pi(p | y) \propto \pi(y | p)\pi(p) \propto p_{11}^{y_{11}} p_{12}^{y_{12}} p_{13}^{y_{13}} p_{21}^{y_{21}} p_{22}^{y_{22}} p_{23}^{y_{23}},$$

and the posterior is thus the distribution

$$\text{Dirichlet}(y_{11} + 1, y_{12} + 1, y_{13} + 1, y_{21} + 1, y_{22} + 1, y_{23} + 1).$$

For the prior predictive distribution, we get

$$\begin{aligned}
\pi(y) &= \frac{\pi(y | p)\pi(p)}{\pi(p | y)} \\
&= \frac{\text{Multinomial}(y | p)}{\text{Dirichlet}(p | y_{11} + 1, y_{12} + 1, \dots, y_{23} + 1)} \\
&= \frac{\Gamma(y_{..} + 1)}{\Gamma(y_{11} + 1)\Gamma(y_{12} + 1)\dots\Gamma(y_{23} + 1)} p_{11}^{y_{11}} \dots p_{23}^{y_{23}} \\
&= \frac{\Gamma(y_{..} + 6)}{\Gamma(y_{11} + 1)\Gamma(y_{12} + 1)\dots\Gamma(y_{23} + 1)} p_{11}^{y_{11}} \dots p_{23}^{y_{23}} \\
&= \frac{\Gamma(y_{..} + 1)}{\Gamma(y_{..} + 6)}.
\end{aligned}$$

(b) Again, we have that $\pi(\theta) = 1$, so

$$\begin{aligned}
\pi(\theta | y) &\propto \pi(y | \theta)\pi(\theta) \\
&\propto (r_1 s_1)^{y_{11}} (r_1 s_2)^{y_{12}} (r_1 s_3)^{y_{13}} (r_2 s_1)^{y_{21}} (r_2 s_2)^{y_{22}} (r_2 s_3)^{y_{23}} \\
&\propto r_1^{y_{1.}} r_2^{y_{2.}} s_1^{y_{.1}} s_2^{y_{.2}} s_3^{y_{.3}},
\end{aligned}$$

so that the posterior is the product of the $\text{Dirichlet}(y_{1.} + 1, y_{2.} + 1)$ distribution and the $\text{Dirichlet}(y_{.1} + 1, y_{.2} + 1, y_{.3} + 1)$ distribution.

For the prior predictive distribution, we get

$$\begin{aligned}
\pi(y) &= \frac{\pi(y | \theta)\pi(\theta)}{\pi(\theta | y)} \\
&= \frac{\text{Multinomial}(y | \theta)}{\text{Dirichlet}(\theta | y_{1.} + 1, y_{2.} + 1, \text{Dirichlet}(\theta | y_{1.} + 1, y_{2.} + 1, y_{3.} + 1))} \\
&= \frac{\frac{\Gamma(y_{..}+1)}{\Gamma(y_{11}+1)\Gamma(y_{12}+1)\dots\Gamma(y_{23}+1)} r_1^{y_1} r_2^{y_2} s_1^{y_{.1}} s_2^{y_{.2}} s_3^{y_{.3}}}{\frac{\Gamma(y_{..}+2)}{\Gamma(y_{1.}+1)\Gamma(y_{2.}+1)} r_1^{y_1} r_2^{y_2} \frac{\Gamma(y_{..}+3)}{\Gamma(y_{.1}+1)\Gamma(y_{.2}+1)\Gamma(y_{.3}+1)} s_1^{y_{.1}} s_2^{y_{.2}} s_3^{y_{.3}}}} \\
&= \frac{\Gamma(y_{..} + 1)\Gamma(y_{1.} + 1)\Gamma(y_{2.} + 1)\Gamma(y_{.1} + 1)\Gamma(y_{.2} + 1)\Gamma(y_{.3} + 1)}{\Gamma(y_{..} + 2)\Gamma(y_{..} + 3)\Gamma(y_{11} + 1)\Gamma(y_{12} + 1) \dots \Gamma(y_{23} + 1)}.
\end{aligned}$$

(c) For the Bayes Factor, we get

$$\begin{aligned}
\text{BF} &= \frac{\pi(y | M_D)}{\pi(y | M_I)} \\
&= \frac{\Gamma(y_{..} + 2)\Gamma(y_{..} + 3)\Gamma(y_{11} + 1)\Gamma(y_{12} + 1) \dots \Gamma(y_{23} + 1)}{\Gamma(y_{..} + 6)\Gamma(y_{1.} + 1)\Gamma(y_{2.} + 1)\Gamma(y_{.1} + 1)\Gamma(y_{.2} + 1)\Gamma(y_{.3} + 1)}
\end{aligned}$$

(d) For the given data, we get

$$\begin{aligned}
\text{BF} &= \frac{\Gamma(6)\Gamma(7)\Gamma(1)\Gamma(2)\Gamma(1)\Gamma(2)\Gamma(1)\Gamma(3)}{\Gamma(10)\Gamma(2)\Gamma(4)\Gamma(2)\Gamma(2)\Gamma(3)} \\
&= \frac{4 \cdot 5}{7 \cdot 8 \cdot 9} = \frac{5}{136} = \frac{1}{27.2}
\end{aligned}$$