

# MSG800/MVE170 Basic Stochastic Processes

Written exam Monday 16 December 2013 2–6 am

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AIDS: Either two A4-sheets (4 pages) of hand-written notes (xerox-copies and/or computer print-outs are not allowed) or Beta (but not both these aids).

GRADES: 12 points for grades 3 and G, 18 points for grade 4, 21 points for grade VG and 24 points for grade 5, respectively.

MOTIVATIONS: All answers/solutions must be motivated.

GOOD LUCK!

**Task 1.** Consider the discrete time random process  $\{X_n, n \geq 1\}$  given by  $X_n = \cos(nU)$  for  $n \geq 1$ , where  $U$  is a random variable that is uniformly distributed over the interval  $[-\pi, \pi]$ . Show that the process  $\{X_n, n \geq 1\}$  is wide-sense stationary. (Hint: The formula  $\cos(x)\cos(y) = \frac{1}{2}\cos(x+y) + \frac{1}{2}\cos(y-x)$  can be useful.) **(5 points)**

**Task 2.** The one and same zero-mean wide sense stationary random process  $\{X(t), t \in \mathbb{R}\}$  with autocorrelation function  $R_X(s, t) = e^{-3|t-s|}$  is input to two different continuous-time linear time-invariant system with outputs  $\{Y_1(t), t \in \mathbb{R}\}$  and  $\{Y_2(t), t \in \mathbb{R}\}$ , respectively, and impulse responses  $h_1(t) = e^{-t}$  for  $t \geq 0$ ,  $h_1(t) = 0$  for  $t < 0$  and  $h_2(t) = e^{-2t}$  for  $t \geq 0$ ,  $h_2(t) = 0$  for  $t < 0$ , respectively. Find  $E(Y_1(t)Y_2(t))$ . **(5 points)**

**Task 3.** Let  $\Phi(x) = \int_{-\infty}^x \frac{1}{\sqrt{2\pi}} e^{-y^2/2} dy$  and  $\phi(x) = \Phi'(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}$  be the standard normal (zero-mean and unit-variance) cumulative probability distribution function and the standard normal probability density function, respectively. Find a random process  $\{X(t), t \in \mathbb{R}\}$  such that the following calculation is valid:

$$\begin{aligned} P(X(1) \leq 0, X(2) \leq 0) &= P(X(1) \leq 0, X(2) - \frac{1}{\sqrt{2}} X(1) \leq -\frac{1}{\sqrt{2}} X(1)) \\ &= \int_{-\infty}^0 P(X(2) - \frac{1}{\sqrt{2}} X(1) \leq -\frac{1}{\sqrt{2}} x) \phi(x) dx \\ &= \int_{-\infty}^0 \Phi(-x) \phi(x) dx \\ &= \int_0^{\infty} \Phi(x) \phi(x) dx \\ &= \left[ \frac{\Phi(x)^2}{2} \right]_0^{\infty} \\ &= \frac{3}{8} \end{aligned} \quad \textbf{(5 points)}$$

Continuation on next page!

**Task 4.** Let  $\{W(t), t \geq 0\}$  be a Wiener process with  $E(W(1)^2) = 1$ . Show that the random process  $\{\int_0^t W(u) du - \frac{1}{3}W(t)^3, t \geq 0\}$  is a martingale with respect to the filtration  $\mathcal{F}_s$  containing information of all values  $\{W(u), u \in [0, s]\}$  of the Wiener process up to time  $s$ . **(5 points)**

**Task 5.** Calculate the limit  $\lim_{s,t \rightarrow \infty} R_X(s, s+t) = \lim_{s,t \rightarrow \infty} E(X(s)X(s+t))$  for a continuous time Markov chain  $\{X(t); t \geq 0\}$  with state space (possible values)  $S$  and generator  $G$  given by

$$S = \{0, 1\} \quad \text{and} \quad G = \begin{pmatrix} -\alpha & \alpha \\ \beta & -\beta \end{pmatrix},$$

respectively, where  $\alpha, \beta > 0$  are given constants. **(5 points)**

**Task 6.** Let  $X(t)$  denote the total number of customers at time  $t \geq 0$  in an M/M/2/4 queuing system in steady-state (/started according to its stationary distribution) with Poisson arrival process with rate  $\lambda = 1$  and with exponentially distributed service times with mean 1. Describe a numerical procedure (/simulation procedure) that gives an approximative numerical value for the probability  $P(\max_{0 \leq t \leq 10} X(t) = 4)$  that the queuing system gets full during the first 10 time units. **(5 points)**

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### Solutions to written exam Monday 16 December

**Task 1.** Clearly  $E(X_n) = 0$  and  $E(X_n X_{n+m}) = \frac{1}{2} E(\cos((2n+m)U)) + \frac{1}{2} E(\cos(mU)) = 0 + \frac{1}{2} \delta(m)$  do not depend on  $n$  so that  $X_n$  is wide sense stationary.

**Task 2.** We have  $E(Y_1(t)Y_2(t)) = E\left(\int_{-\infty}^{\infty} h_1(u)X(t-u) du \int_{-\infty}^{\infty} h_2(v)X(t-v) dv\right) = E\left(\int_0^{\infty} e^{-u}X(t-u) du \int_0^{\infty} e^{-2v}X(t-v) dv\right) = \int_0^{\infty} \int_0^{\infty} e^{-u}e^{-2v}E(X(t-u)X(t-v)) dudv = \int_0^{\infty} \int_0^{\infty} e^{-u}e^{-2v}e^{-3|u-v|} dudv = \int_{u=0}^{u=\infty} \int_{v=u}^{v=\infty} e^{-u}e^{-2v}e^{-3(v-u)} dvdu + \int_{v=0}^{v=\infty} \int_{u=v}^{u=\infty} e^{-u}e^{-2v}e^{-3(u-v)} dudv = \int_{u=0}^{u=\infty} e^{2u} \int_{v=u}^{v=\infty} e^{-5v} dvdu + \int_{v=0}^{v=\infty} e^v \int_{u=v}^{u=\infty} e^{-4u} dudv = \int_{u=0}^{u=\infty} \frac{1}{5} e^{-3u} du + \int_{v=0}^{v=\infty} \frac{1}{4} e^{-3v} dv = \frac{1}{15} + \frac{1}{12} = \frac{3}{20}$ .

**Task 3.** A zero-mean unit-variance Gaussian process with  $E(X(1)X(2)) = \frac{1}{\sqrt{2}}$ , because for such a process  $X(1)$  and  $X(2) - \frac{1}{\sqrt{2}}X(1)$  are independent and  $P(X(2) - \frac{1}{\sqrt{2}}X(1) \leq -\frac{1}{\sqrt{2}}x) = \Phi(-x)$  since  $X(2) - \frac{1}{\sqrt{2}}X(1)$  is a Gaussian random variable with variance  $\frac{1}{2}$ .

**Task 4.** We have

$$\begin{aligned} & E\left(\int_0^t W(u) du - \frac{1}{3}W(t)^3 \middle| \mathcal{F}_s\right) \\ &= E\left(\int_s^t (W(u) - W(s)) du \middle| \mathcal{F}_s\right) + E((t-s)W(s) \middle| \mathcal{F}_s) + E\left(\int_0^s W(u) du \middle| \mathcal{F}_s\right) \\ &\quad - \frac{1}{3}E((W(t) - W(s))^3 \middle| \mathcal{F}_s) - E((W(t) - W(s))^2 W(s) \middle| \mathcal{F}_s) \\ &\quad - E((W(t) - W(s))W(s)^2 \middle| \mathcal{F}_s) - \frac{1}{3}E(W(s)^3 \middle| \mathcal{F}_s) \\ &= E\left(\int_s^t (W(u) - W(s)) du\right) + (t-s)W(s) + \int_0^s W(u) du \\ &\quad - \frac{1}{3}E((W(t) - W(s))^3) - E((W(t) - W(s))^2 W(s)) \\ &\quad - E(W(t) - W(s))W(s)^2 - \frac{1}{3}W(s)^3 \\ &= 0 + (t-s)W(s) + \int_0^s W(u) du - 0 - (t-s)W(s) - 0 - \frac{1}{3}W(s)^3 \\ &= \int_0^s W(u) du - \frac{1}{3}W(s)^3 \quad \text{for } 0 \leq s \leq t. \end{aligned}$$

**Task 5.** The chain is irreducible with stationary distribution  $\pi = \left(\frac{\beta}{\alpha+\beta} \quad \frac{\alpha}{\alpha+\beta}\right)$  (as this gives  $\pi G = 0$ ). Noting that  $X(s)X(t) = 1$  when both  $X(s)$  and  $X(t)$  are 1 while  $X(s)X(t) = 0$  otherwise it follows that  $E(X(s)X(t)) = (\mu^{(s)})_1 p_{11}(t) = (\mu^{(0)}P_s)_1 p_{11}(t) = ((\mu^{(0)})_0 p_{01}(s) + (\mu^{(0)})_1 p_{11}(s)) p_{11}(t) \rightarrow ((\mu^{(0)})_0 \pi_1 + (\mu^{(0)})_1 \pi_1) \pi_1 = \pi_1^2 = \alpha^2/(\alpha+\beta)^2$  as  $s, t \rightarrow \infty$ .

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Clear[slump, start, xt, tid];
In[3]:= N[Mean[Table[{tid = 0, xt = {slump = Random[UniformDistribution[{0, 1}]}], If[slump < 8/23, start = 0],
  If[8/23 ≤ slump < 16/23, start = 1], If[16/23 ≤ slump < 20/23, start = 2],
  If[20/23 ≤ slump < 22/23, start = 3], If[22/23 ≤ slump, start = 4], start}][[7]],
While[tid < 10 && xt < 7/2, If[xt < 1/2, tid = tid + Random[ExponentialDistribution[1]]; xt = 1,
  If[1/2 < xt < 3/2,
    tid = tid + Random[ExponentialDistribution[2]]; slump = Random[UniformDistribution[{0, 1}]];
    If[slump ≤ 1/2, xt = xt + 1, xt = xt - 1], tid = tid + Random[ExponentialDistribution[3]];
    slump = Random[UniformDistribution[{0, 1}]]; If[slump ≤ 1/3, xt = xt + 1, xt = xt - 1]]],
  If[xt > 7/2 && tid < 10, 1, 0]][[4]], {i, 1, 100 000}]]]
Out[3]= 0.46679

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