

**Solutions for the exam for Matematisk statistik och diskret matematik (MVE050/MSG810). Statistik för fysiker (MSG820).
December 15, 2012.**

1. (3p) The joint distribution of the discrete random variables X and Y is given by the density table:

$x \backslash y$	0	1
1	0.4	0.2
2	0.1	0.3

- a) Find $\mathbf{E}[X]$, $\mathbf{E}[Y]$, $\mathbf{E}[XY]$
 b) Find $\mathbf{Cov}(X, Y)$
 c) Find $\mathbf{E}[5X - Y + 17]$

Solution:

a)

$$\mathbf{E}[X] = 1 * (0.4 + 0.2) + 2 * (0.1 + 0.3) = 1.4,$$

$$\mathbf{E}[Y] = 0 * (0.4 + 0.1) + 1 * (0.2 + 0.3) = 0.5,$$

$$\mathbf{E}[XY] = 0.4 * 0 * 1 + 0.2 * 1 * 1 + 0.1 * 0 * 2 + 0.3 * 1 * 2 = 0.8$$

b) $\mathbf{Cov}(X, Y) = \mathbf{E}[XY] - \mathbf{E}[X]\mathbf{E}[Y] = 0.8 - 1.4 * 0.5 = 0.1$

c) $\mathbf{E}[5X - Y + 17] = 5 * \mathbf{E}[X] - \mathbf{E}[Y] + 17 = 5 * 1.4 - 0.5 + 17 = 23.5$

2. (3p) Draw a card at random from a standard 52-card deck. Denote $A = \{\text{draw a nine}\}$, $B = \{\text{draw a diamond}\}$, $C = \{\text{draw a number}\}$.

- a) Are the events B and C independent? A and C ? Why?
 b) Find $\mathbf{P}(A \cap B|C)$.
 c) Remove one card (the Ace of spades) from the deck. Are A and B independent now? Why?

Solution:

a) B and C are independent:

$$\mathbf{P}(B \cap C) = \mathbf{P}(\text{draw a number which is a diamond}) = 9/52,$$

$$\mathbf{P}(B) = 1/4, \mathbf{P}(C) = 9/13,$$

so $\mathbf{P}(B \cap C) = \mathbf{P}(B)\mathbf{P}(C)$ – a definition of independent events.

Similarly, show that A and C are not independent:

$$\mathbf{P}(A \cap C) = 1/52 \neq 1/52 * 9/13 = \mathbf{P}(A) * \mathbf{P}(C).$$

b)

$$\mathbf{P}(A \cap B|C) = \frac{\mathbf{P}(A \cap B)}{\mathbf{P}(C)} = \frac{1/52}{9/13} = \frac{1}{36}.$$

c) No, if we remove one card, A and B are not independent anymore. Now:

$$\mathbf{P}(A \cap B) = 1/51 \neq 4/51 * 13/51 = \mathbf{P}(A) * \mathbf{P}(B).$$

3. (2p) The continuous random variable X has a density $f_X(x)$:

$$f_X(x) = \begin{cases} \frac{x}{a}, & 20 < x < 40, \\ 0, & \text{otherwise.} \end{cases}$$

a) Find the value of a so that $f_X(x)$ becomes a proper density function.

b) Using the obtained value of a , find the probability $\mathbf{P}(X < 25)$.

Solution:

a) For $f_X(x)$ to be a proper density function, it must be non-negative and the total area under the curve must be equal to 1:

$$\int_{-\infty}^{+\infty} f_X(x) dx = \int_{20}^{40} \frac{x}{a} dx = \frac{x^2}{2a} \Big|_{20}^{40} = \frac{1600 - 400}{2a} = \frac{600}{a} = 1,$$

so $a = 600$.

b)

$$\begin{aligned} \mathbf{P}(X < 25) &= \int_{-\infty}^{25} f_X(x) dx = \int_{20}^{25} \frac{x}{600} dx = \frac{x^2}{1200} \Big|_{20}^{25} \\ &= \frac{625 - 400}{1200} = \frac{3}{16} = 0.1875 \end{aligned}$$

4. (4p) Assume the total proportion p of red-haired people in Sweden is 0.015. The dentist examines $n = 10$ people every day, which you can assume to be independent of each other.

a) What is the probability for the dentist to meet at least 2 red-haired individuals during the same day?

b) What is the probability to examine at least 1 red-haired person every day during the whole working week (5 days)?

c) What is the expected number of days a new dentist has to work before meeting the first red-haired person?

Solution:

a) Denote by X the total number of red-haired people a doctor examines during one day. X can be thought of as a Binomially($n = 10$, $p = 0.015$) distributed random variable. We are interested in the probability:

$$\begin{aligned} \mathbf{P}(X \geq 2) &= 1 - \mathbf{P}(X \leq 1) = 1 - \mathbf{P}(X = 0) - \mathbf{P}(X = 1) \\ &= 1 - \binom{10}{0} (0.015)^0 (0.985)^{10} - \binom{10}{1} (0.015)^1 (0.985)^9 = 0.009 \end{aligned}$$

b) First, find the probability to meet at least one red-haired person during one day.

$$\mathbf{P}(X \geq 1) = 1 - \mathbf{P}(X = 0) = 1 - \binom{10}{0} (0.015)^0 (0.985)^{10} = 1 - 0.860 = 0.140$$

Now, assuming the different days being independent of each other, we can argue:

$$\begin{aligned} & \mathbf{P}(\text{meet at least 1 on every one of 5 working days}) \\ &= \mathbf{P}\left(\bigcap_{i=1}^5 \{\text{meet at least 1 red-haired person on } i\text{-th day}\}\right) \\ &= (\mathbf{P}(X \geq 1))^5 = (0.140)^5 \approx 0.0000538 \end{aligned}$$

c) There is several ways to approach this problem, here is one of them. Denote by Y the number of the first day the doctor meets a red-haired patient. We can see the sequence of days as a sequence of independent Bernoulli(0.140) trials, since the probability to meet at least one red-haired patient during one day is 0.140 (from part b). Therefore, Y has a Geometric distribution with parameter 0.140, hence

$$\mathbf{E}Y = \frac{1}{0.140} = 7.14$$

5. (3p) Bad thoughts are popping up in Peter's head according to a Poisson process with intensity $\lambda = 2$ [tph] (thoughts per hour). When bad thoughts are there, Peter becomes uneasy and can not eat.

- a) It takes Peter 40 minutes to eat his lunch. What's the probability for him to eat the lunch in peace, without any bad thoughts?
- b) If Peter wakes up 6:30 and has a breakfast at 7, what's the probability for him to get at least 3 bad thoughts before breakfast?

Solution:

a) The amount X of bad thoughts that come to Peter during lunch has a Poisson distribution with parameter λt , where $\lambda = 2$ and $t = 2/3$. We are interested in the probability

$$\mathbf{P}(X = 0) = e^{-\lambda t} \frac{(\lambda t)^0}{0!} = e^{-4/3} \approx 0.264$$

b) Again, the amount of bad thoughts $X \sim \text{Pois}(\lambda t)$. Now t is 0.5, $\lambda t = 1$, and we want the probability

$$\begin{aligned} \mathbf{P}(X \geq 3) &= 1 - \mathbf{P}(X \leq 2) = 1 - e^{-1} \frac{1^0}{0!} - e^{-1} \frac{1^1}{1!} - e^{-1} \frac{1^2}{2!} \\ &= 1 - e^{-1}(1 + 1 + 0.5) = 1 - 2.5e^{-1} \approx 0.0803 \end{aligned}$$

6. (3p) Find a generating function $A(x)$ of the sequence (a_n) , defined by the following recursion:

$$\begin{cases} a_0 = 2, \\ a_1 = 2, \\ a_n = a_{n-1} + a_{n-2} - 1, \quad n = 2, 3, \dots \end{cases}$$

Note: you don't have to find the (a_n) itself, just the generating function.

Solution:

Start with writing out the recursion for $n = 2, 3, \dots$ and multiplying with the corresponding power of x :

$$\begin{array}{ll} a_2 = a_1 + a_0 - 1, & *x^2 & a_2x^2 = xa_1x + x^2a_0 - x^2 * 1 \\ a_3 = a_2 + a_1 - 1, & *x^3 & a_3x^3 = xa_2x^2 + x^2a_1x - x^2 * x \\ a_4 = a_3 + a_2 - 1, & *x^4 & a_4x^4 = xa_3x^3 + x^2a_2x^2 - x^2 * x^2 \\ \dots & & \dots \end{array}$$

Sum up:

$$\begin{aligned} a_2x^2 + a_3x^3 + a_4x^4 + \dots &= x(a_1x + a_2x^2 + a_3x^3 + \dots) \\ &+ x^2(a_0 + a_1x + a_2x^2 + \dots) - x^2(1 + x + x^2 + \dots) \end{aligned}$$

Now, remembering that $\sum_{n=0}^{\infty} a_nx^n = A(x)$ and $\sum_{n=0}^{\infty} x^n = \frac{1}{1-x}$, we get:

$$A(x) - a_1x - a_0 = x(A(x) - a_0) + x^2A(x) - x^2 \frac{1}{1-x},$$

or, putting in the values for a_0, a_1 :

$$A(x) - 2x - 2 = x(A(x) - 2) + x^2A(x) - \frac{x^2}{1-x}.$$

Combine the terms with $A(x)$ on the left and the rest on the right:

$$A(x)(1 - x - x^2) = 2 - \frac{x^2}{1-x} = \frac{2 - 2x - x^2}{1-x},$$

so

$$A(x) = \frac{2 - 2x - x^2}{(1-x)(1-x-x^2)}.$$

7. (4p) Stefan is on a quest to find out the average IQ of a seagull. Omitting details about how he measures the IQ of his involuntary subjects, we present his data:

7.5 7.0 8.5 4.0 6.0 7.0 5.5

- Help Stefan construct the 95% confidence interval for μ – the actual average IQ of a seagull.
- What's the length of your interval? Based on your estimate for the actual variance, how many seagulls does Stefan have to catch for his 95% confidence interval to be of length <1 ?

Solution:

- a) We assume Stefan's data to be $\text{Normal}(\mu, \sigma)$ with some unknown parameters. That means, the quantity $T = \frac{\bar{X} - \mu}{s/\sqrt{n}}$ has t -distribution with $n - 1$ degrees of freedom (in our case, $n = 7$). That fact leads us to the formulas for the left and right border of the two-sided 95% CI:

$$L = \bar{X} - t_{6,0.025} \frac{s}{\sqrt{7}}, R = \bar{X} + t_{6,0.025} \frac{s}{\sqrt{7}}$$

Looking up $t_{6,0.025} = 2.447$ in the table and having $\bar{X} = 6.5$ and $s = 1.472$ found from the data, we obtain the confidence interval:

$$[5.1386, 7.8606]$$

- b) The length $l = R - L$ of the obtained confidence interval is 2.722. However, we can decrease that by making more observations:

$$l = R - L = 2t_{6,0.025} \frac{s}{\sqrt{n}} = 2 * 2.447 \frac{1.472}{\sqrt{n}} = \frac{7.204}{\sqrt{n}},$$

so for that quantity to be less than one (assuming the sample standard deviation stays the same) we need n to be greater than $(7.204)^2 = 51.9$, so $n = 52$ should be enough.

Note: we are overestimating a bit, since we are not taking into account how $t_{n-1,0.025}$ decreases when we increase n .

8. (5p) Tim started working as a ticket controller for a bus company in the beginning of November. Before this job, he had a feeling that around 20% of people under 21 years old usually did not have tickets. During first month of his job, out of $n = 132$ such young people he met, 17 were without tickets. Is there a statistically significant reason to update his old beliefs? Test the hypothesis $H_0 : p = 0.2$ against the appropriate alternative. Decide if you need a two- or one-sided test. What's the p -value for the test? Can you reject H_0 on a significance level $\alpha = 0.01$?

Solution:

We are going to conduct a significance test of a hypothesis $H_0 : p = 0.2$ against the alternative $H_1 : p < 0.2$ to see if the data collected by Tim is significant enough to update his old belief (20% of young people don't buy tickets) to a new one (less than 20% of young people don't buy tickets). The test statistic we are going to use is $Z = \frac{\hat{p} - p_0}{\sqrt{p_0(1-p_0)/n}}$, here $p_0 = 0.2$ is the default value of p and \hat{p} is the point estimate for the proportion, obtained from data. Due to the Central Limit Theorem, under H_0 , Z has a distribution very close to $\text{Normal}(0, 1)$. Find the value of Z . From the data, $\hat{p} = 17/132 \approx 0.129$, so

$$Z = \frac{0.129 - 0.2}{\sqrt{0.2 * 0.8/132}} \approx -2.04$$

So, the p -value for the test, obtained from the tables for cdf of a $\text{Normal}(0,1)$ random variable, is

$$p\text{-value} = 0.0207$$

The obtained p -value is small, but the result is not significant on the level $\alpha = 0.01$, hence we cannot reject H_0 on that significance level.

Note, it is also possible to test H_0 against the alternative $H_1 : p \neq 0.2$, in that case the p -value would double, to become 0.0414.

9. (3p) Barbara is playing roulette in a casino, always putting in the minimal bet of 2 cents. Let us denote by X_i the change in her capital after i -th game. One can assume all of the games to be independent of each other, played by the same rules (thus X_i 's are distributed identically to some random variable X). The expected gain after one game is negative: $\mu_X = -0.06$ cents, the standard deviation is $\sigma_X = 2$ cents. Barbara has 2 dollars, her plan is to play $n = 100$ games and then quit, no matter what.

- a) Find the expected value of Barbara's total gain $S_{100} = X_1 + X_2 + \dots + X_{100}$.
- b) Find an approximate probability for her total gain S_{100} to be positive. (Hint: use the Central Limit Theorem)

Solution:

- a) The expectation of the total gain:

$$\begin{aligned}\mathbf{E}S_{100} &= \mathbf{E}[X_1 + X_2 + \dots + X_{100}] = \mathbf{E}[X_1] + \mathbf{E}[X_2] + \dots + \mathbf{E}[X_{100}] \\ &= -0.06 - 0.06 - \dots - 0.06 = -100 * 0.06 = -6\end{aligned}$$

- b) By CLT, the distribution of $Z = (S_{100} - 100\mu_X)/\sqrt{100}\sigma_X$ can be approximated with Normal(0, 1), so

$$\mathbf{P}(S_{100} > 0) = \mathbf{P}(S_{100} - 100\mu_X / 10\sigma_X > -100\mu_X / 10\sigma_X) = \mathbf{P}(Z > -10\mu_X / \sigma_X) = \mathbf{P}(Z > 0.3)$$

The latter can be found from the tables for Normal distribution to be 0.382