

May 4, 2004

**Solutions: Chapter 8**

Problems 3, 4, 28, 45, 47

**Problem 3.** Number of yeast cells on a square  $X$ . Test the Poisson model  $X \in \text{Pois}(\lambda)$ .

Concentration 1.  $\bar{X} = 0.6825$ ,  $\overline{X^2} = 1.2775$ ,  $s^2 = 0.8137$ ,  $s = 0.9021$ ,  $s_{\bar{X}} = 0.0451$ . Approximate 95% CI for  $\mu$ :  $0.6825 \pm 0.0884$ .

Pearson's chi-square test based on  $\hat{\lambda} = 0.6825$ :

$x$	0	1	2	3	4+	Total
Observed	213	128	37	18	4	400
Expected	202.14	137.96	47.08	10.71	2.12	400

Observed test statistic  $X^2 = 10.12$ ,  $\text{df} = 5 - 1 - 1 = 3$ ,  $P < 0.025$ . Reject the model.

Concentration 2.  $\bar{X} = 1.3225$ ,  $\overline{X^2} = 3.0325$ ,  $s = 1.1345$ ,  $s_{\bar{X}} = 0.0567$ . Approximate 95% CI for  $\mu$ :  $1.3225 \pm 0.1112$ .

Pearson's chi-square test: observed test statistic  $X^2 = 3.16$ ,  $\text{df} = 4$ ,  $P > 0.10$ . Accept the model.

Concentration 3.  $\bar{X} = 1.8000$ ,  $s = 1.1408$ ,  $s_{\bar{X}} = 0.0701$ . Approximate 95% CI for  $\mu$ :  $1.8000 \pm 0.1374$ .

Pearson's chi-square test: observed test statistic  $X^2 = 7.79$ ,  $\text{df} = 5$ ,  $P > 0.10$ . Accept the model.

Concentration 4.  $n = 410$ ,  $\bar{X} = 4.5659$ ,  $s^2 = 4.8820$ ,  $s_{\bar{X}} = 0.1091$ . Approximate 95% CI for  $\mu$ :  $4.566 \pm 0.214$ .

Pearson's chi-square test: observed test statistic  $X^2 = 13.17$ ,  $\text{df} = 10$ ,  $P > 0.10$ . Accept the model.

**Problem 4.** Likelihood function of  $X \in \text{Bin}(n, p)$  for a given  $n$  and  $X = x$  is

$$L(p) = \binom{n}{x} p^x (1-p)^{n-x} \propto p^x (1-p)^{n-x}.$$

a) Minimize  $\ln p^x (1-p)^{n-x} = x \ln p + (n-x) \ln(1-p)$ . Since

$$\frac{\partial}{\partial p} (x \ln p + (n-x) \ln(1-p)) = \frac{x}{p} - \frac{n-x}{1-p}$$

solve  $\frac{x}{p} = \frac{n-x}{1-p}$  to find MLE  $\hat{p} = \frac{x}{n}$ .

b) Cramer-Rao: if  $\tilde{p}$  is unbiased, then  $\text{Var}(\tilde{p}) \geq (kI(p))^{-1} = 1/I(p)$  since the sample size is  $k = 1$ . Here

$$I(p) = -E\left(\frac{d^2}{dp^2} \ln f(X|p)\right) \text{ and } f(X|p) = \binom{n}{X} p^X (1-p)^{n-X}.$$

Compute

$$\ln f(X|p) = \text{const} + X \ln p + (n-X) \ln(1-p), \quad \frac{d}{dp} \ln f(X|p) = \frac{X}{p} - \frac{n-X}{1-p},$$

$$\frac{d^2}{dp^2} \ln f(X|p) = -\frac{X}{p^2} - \frac{n-X}{(1-p)^2}, \text{ and } I(p) = E\left(\frac{X}{p^2} + \frac{n-X}{(1-p)^2}\right) = \frac{n}{p(1-p)}.$$

Thus  $\text{Var}(\hat{p}) \geq 1/I(p) = \frac{p(1-p)}{n} = \text{Var}(\hat{p})$ .

c) Plot  $L(p) = 252p^5(1-p)^5$ .

**Problem 28.**

a)  $\bar{X} = 3.6109$  and  $s^2 = 3.4181$ . Here  $n = 16$  and  $s_{\bar{X}} = 0.4622$ .

b) and c) Exact CI

	90%	95%	99%
$\mu$	$3.61 \pm 0.81$	$3.61 \pm 0.98$	$3.61 \pm 1.36$
$\sigma^2$	(2.05; 7.06)	(1.87; 8.19)	(1.56; 11.15)
$\sigma$	(1.43; 2.66)	(1.37; 2.86)	(1.25; 3.34)

d) Find sample size  $x$  to halve the CI length:

$$t_{15}(\alpha/2) \cdot \frac{s}{\sqrt{16}} = 2 \cdot t_{x-1}(\alpha/2) \cdot \frac{s'}{\sqrt{x}} \text{ implies } x \approx (2 \cdot 4)^2 = 64.$$

Further adjustment for 95% CI:

$$t_{15}(\alpha/2) = 2.13, t_{x-1}(\alpha/2) \approx 2, \text{ therefore } x \approx (2 \cdot 4 \cdot \frac{2}{2.13})^2 = 56.4.$$

**Problem 45.** Uniform distribution  $U(0, \theta)$ :

$$f(x) = \frac{1}{\theta} \text{ for } 0 \leq x \leq \theta, f(x) = 0 \text{ otherwise.}$$

a)  $\mu = \theta/2, \tilde{\theta} = 2\bar{X}, E(\tilde{\theta}) = \theta, \text{Var}(\tilde{\theta}) = \frac{4\sigma^2}{n} = \frac{\theta^2}{3n}$ .

b)  $L(\theta) = \frac{1}{\theta^n}$  for  $\theta \geq X_{(n)}, L(\theta) = 0$  otherwise. MLE  $\hat{\theta} = X_{(n)} = \max(X_1, \dots, X_n)$ .

c) Sampling distribution of the MLE  $\hat{\theta} = X_{(n)}$ :

$$P(X_{(n)} \leq x) = (\frac{x}{\theta})^n \text{ with pdf } f_{\hat{\theta}} = \frac{n}{\theta^n} \cdot x^{n-1}, 0 \leq x \leq \theta$$

$$\text{biased } E(\hat{\theta}) = \frac{n}{n+1}\theta, E(\hat{\theta}^2) = \frac{n}{n+2}\theta^2, \text{Var}(\hat{\theta}) = \frac{\theta^2}{(n+1)^2(n+2)}.$$

Mean square errors:

$$\text{MSE}(\hat{\theta}) = (-\frac{\theta}{n+1})^2 + \frac{\theta^2}{(n+1)^2(n+2)} = \frac{n+3}{n+2} \cdot \frac{\theta^2}{(n+1)^2}$$

$$\text{MSE}(\tilde{\theta}) = \frac{\theta^2}{3n}$$

d) Corrected MLE  $\hat{\theta}_c = \frac{n+1}{n} \cdot X_{(n)}$  with  $E(\hat{\theta}_c) = \theta, \text{Var}(\hat{\theta}_c) = \frac{\theta^2}{n^2(n+2)}$ .

**Problem 47.** Genetic model:  $p_1 = \frac{2+\theta}{4}, p_2 = \frac{1-\theta}{4}, p_3 = \frac{1-\theta}{4}, p_4 = \frac{\theta}{4}$ , where  $0 < \theta < 1$ . In particular, if  $\theta = 0.25$ , then the genes are unlinked and the genotype frequencies are

	Green	White	Total
Starchy	$\frac{9}{16}$	$\frac{3}{16}$	$\frac{3}{4}$
Sugary	$\frac{3}{16}$	$\frac{1}{16}$	$\frac{1}{4}$
Total	$\frac{3}{4}$	$\frac{1}{4}$	1

a) Sample counts  $(X_1, X_2, X_3, X_4) \in \text{Mn}(n, p_1, p_2, p_3, p_4)$  with  $n = 3839$ .  
Likelihood

$$L(\theta) = \binom{n}{x_1, x_2, x_3, x_4} (2 + \theta)^{x_1} (1 - \theta)^{x_2 + x_3} \theta^{x_4} 4^{-n}$$

$$\frac{d}{d\theta} \ln L(\theta) = \frac{x_1}{2 + \theta} - \frac{x_2 + x_3}{1 - \theta} + \frac{x_4}{\theta}.$$

Solve

$$\frac{x_1}{2 + \theta} + \frac{x_4}{\theta} = \frac{x_2 + x_3}{1 - \theta} \text{ or } \theta^2 n + \theta(2x_2 + 2x_3 + x_4 - x_1) - 2x_4 = 0$$

to find MLE  $\hat{\theta} = \frac{-u + \sqrt{u^2 + 8nx_4}}{2n} = 0.0357$ , where  $u = 2x_2 + 2x_3 + x_4 - x_1$ .

Asymptotic variance  $\text{Var}(\hat{\theta}) \approx \frac{1}{I(\hat{\theta})}$ , where  $I(\theta) = -E\left(\frac{d^2}{d\theta^2} \ln f(X_1, X_2, X_3, X_4|\theta)\right)$ .

Since

$$\frac{d^2}{d\theta^2} \ln L(\theta) = -\frac{x_1}{(2 + \theta)^2} - \frac{x_2 + x_3}{(1 - \theta)^2} - \frac{x_4}{\theta^2}$$

$$I(\theta) = \frac{n}{4(2 + \theta)} + \frac{2n}{4(1 - \theta)} + \frac{n}{4\theta} = \frac{n(1 + 2\theta)}{2\theta(2 + \theta)(1 - \theta)} \text{ and } I(\hat{\theta}) = 29345.8$$

we have  $s_{\hat{\theta}} = 0.0058$ .

b)  $0.0357 \pm 1.96 \cdot 0.0058 = 0.0357 \pm 0.0114$

c) Parametric bootstrap using Matlab:

```
p1=0.5089, p2=0.2411, p3=0.2411, p4=0.0089
n=3839; B=1000; b=ones(B,1);
x1=binornd(n,p1,B,1);
x2=binornd(n*b-x1,p2/(1-p1));
x3=binornd(n*b-x1-x2,p3/(1-p1-p2));
x4=n*b-x1-x2-x3;
u=2*x2+2*x3+x4-x1;
t=(-u+sqrt(u.^2+8*n*x4))/(2*n);
std(t)
histfit(t)
```

gives  $\text{std}(t) = 0.0058$ .

d) Two ends of interval covering 95% of the components of the vector t produced by bootstrapping:

$$c1 = \text{prctile}(t, 2.5)$$

$$c2 = \text{prctile}(t, 97.5)$$

are  $c1 = 0.0250$  and  $c2 = 0.0473$ . 95% CI for  $\theta$ :  $(2\hat{\theta} - c_2, 2\hat{\theta} - c_1) = (0.0241, 0.0464)$