

Checking Adequacy of Model

(21)

Assuming regression model is adequate. We can use it to make inferences, see previous pages. Before doing so, we need to check if the assumptions hold.

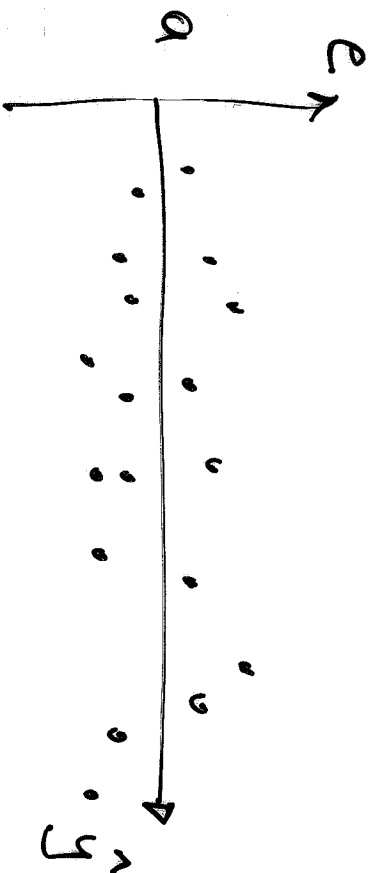
- Are $\varepsilon \sim N(0, \sigma^2)$?

First we compute the residuals, after a, b the estimates of α and β have been estimated.

Then

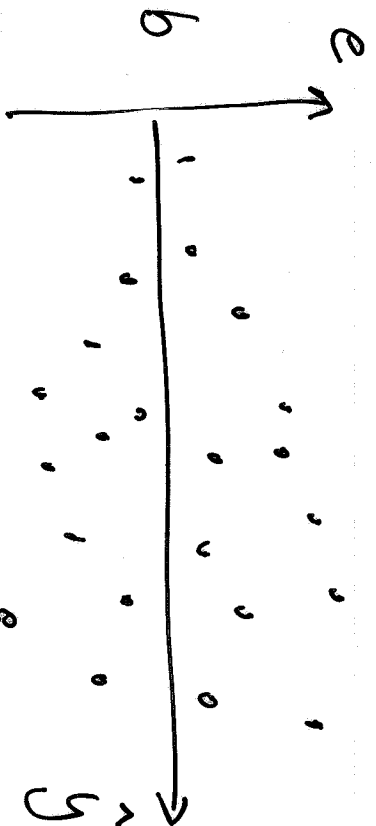
$$e_i = y_i - \hat{y}_i = y_i - (a + bx_i)$$

- Normplot the residuals e_i to check if the normality assumption holds
- Plot residuals vs $\hat{y}_i$ (predicted values)

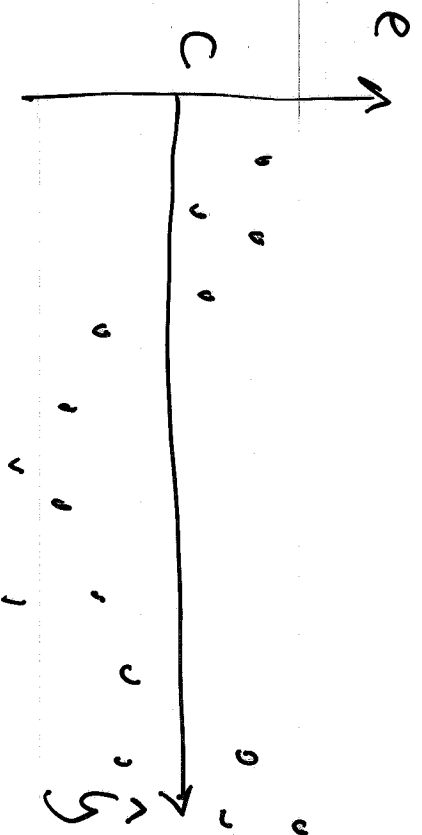


ideal
"constant" band

decrease!



variance increases
with the
response and
a transformation
is needed



linear model
is NOT the
proper model.

- Plot residuals vs time (or order when this applies) to detect any trends over time.

When variance of residuals is not constant, like in Figure b, we need to do a transformation in order to "stabilize" the variance.

It is usual practise then to substitute in the regression

$$Y = \alpha + \beta X + \varepsilon \quad \text{or}$$

$$\hat{Y} = \alpha + b X$$

the independent variable X by a transformed one. Usually $X' = \log X$

$$\text{or } X' = \sqrt{X} \quad \text{or } X' = X^k \quad \text{for some } k,$$

are the most common ones.

Be careful, $\log X$, $\sqrt{X}$ can be used when X takes only positive values!

Correlation

(24)

Up to now, we have assumed that the independent variable $\underline{X}$ was assumed to be known without error. In some cases though both x and y have to be thought of as values of random vars $\underline{X}, Y$.

Example Study relationship between input and output of a wastewater treatment plant. / tensile strength and hardness of aluminium ...

In such cases data $(x_i, y_i), i=1, \dots, n$

are values of a pair of random variables

$\underline{X}$ and $\underline{Y}$ having joint density

$$f_{\underline{X}, \underline{Y}}(x, y).$$

Scatterplot provides visual impression of the relation between x and y .

Chosen of the scatter plot to a straight line can be expressed in terms of correlation coefficient (numerically).

Sample correlation coefficient, r , is defined

$$r = \frac{1}{n-1} \sum_{i=1}^n \left(\frac{x_i - \bar{x}}{s_x} \right) \cdot \left(\frac{y_i - \bar{y}}{s_y} \right) = \frac{s_{xy}}{\sqrt{s_{xx} \cdot s_{yy}}}$$

Where $s_x^2 = \sum_{i=1}^n \frac{(x_i - \bar{x})^2}{n-1} = \frac{s_{xx}}{n-1}$

$$s_y^2 = \sum_{i=1}^n \frac{(y_i - \bar{y})^2}{n-1} = \frac{s_{yy}}{n-1}$$

- When most of the pairs of obs are s.t. either both (x_i, y_i) are above their sample means or both below their sample means the products tend to be large and positive so $r > 0$.

• When one value of the pair tends to be large when the other is small and vice versa, $r < 0$. (26)

• $-1 \leq r \leq 1$.

1) Magnitude of r describes the strength of a linear relation and sign of r indicates the direction

• $r = 1$ - all (x_i, y_i) lie exactly on a straight line having positive slope.

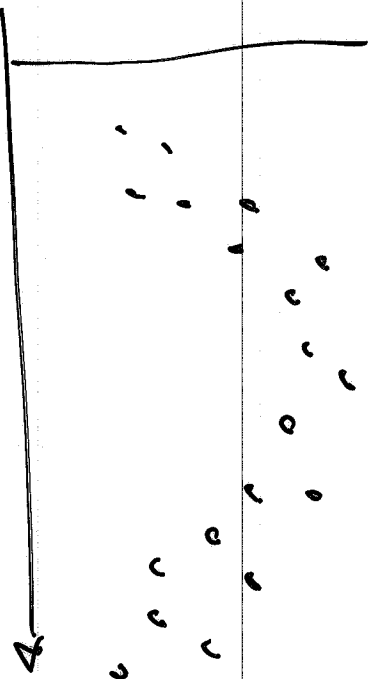
• $r > 0$ - pattern of scatter plot runs from lower left to upper right

• $r < 0$ - pattern of scatter plot runs from upper left to lower right.

• $r = -1$ all pairs (x_i, y_i) lie on a straight line having a negative slope.

• A value of r near $+1$ or -1 describes strong linear relation.

2) A value of r close to $0 \Rightarrow$ linear association is weak but there can be still strong association along for example, a curve



Exploring and interpreting

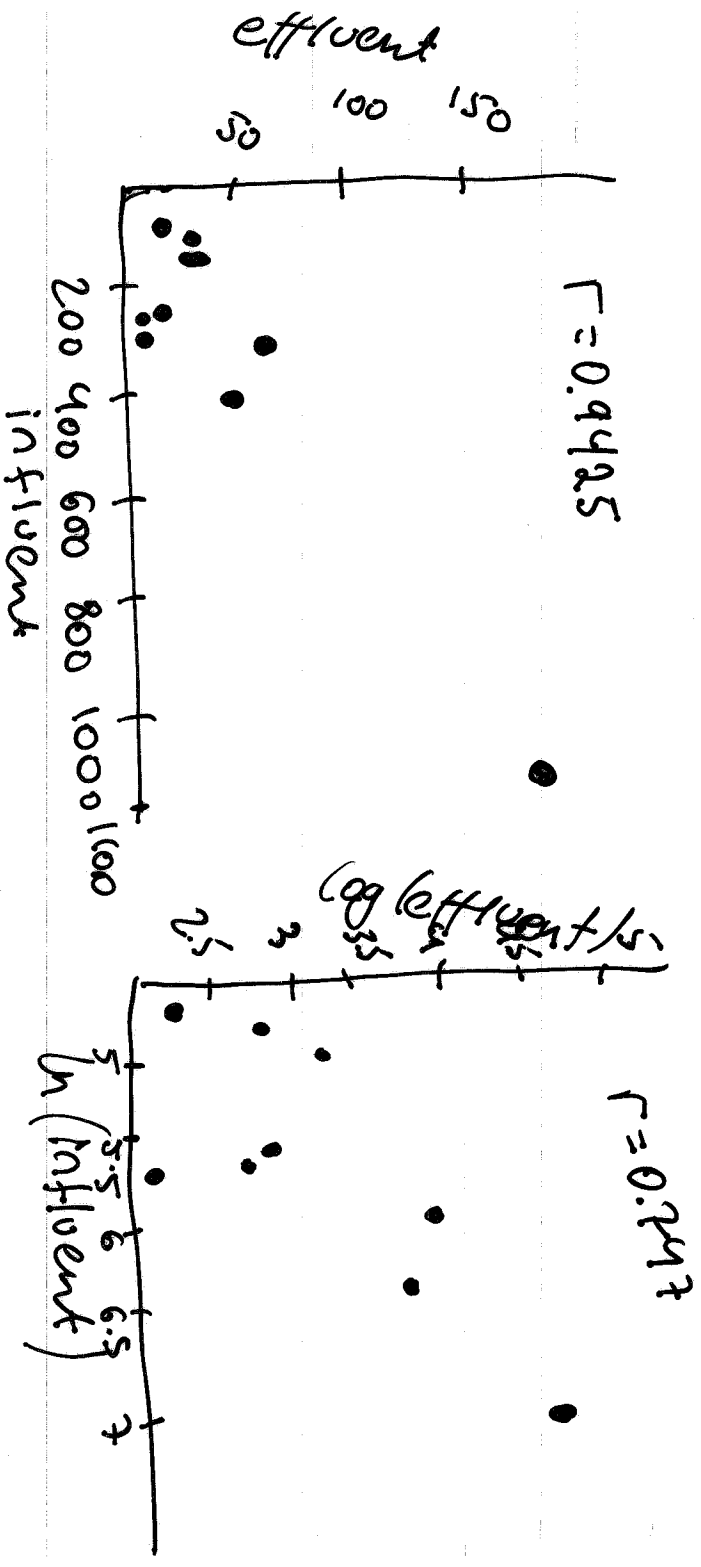
(28)

correlation

Example

Heavy metals can inhibit the biological treatment of waste in municipal treatment plants. Monthly measurements were made at a state-of-the-art treatment plant of the amount of Chromium ($\mu\text{g/l}$) in both the influent and effluent.

Influent	250	200	270	100	300	410	110	130	1100
Effluent	19	10	17	11	70	60	18	30	180



Notice that r is not really appropriate for the original data since the one large obs in the upper right hand corner has too much influence. If the pair (1100, 180) is dropped, r drops to 0.578.

Correlation and Regression

There are two important relationships between r and the least square fit of a straight line.

$$r = \frac{S_{xy}}{\sqrt{S_{xx} \cdot S_{yy}}} = \frac{\sqrt{S_{xx}}}{\sqrt{S_{yy}}} \frac{S_{xy}}{S_{xx}} = \frac{\sqrt{S_{xx}}}{\sqrt{S_{yy}}} \cdot b \Rightarrow$$

r and b have the same sign.

total variation in y (30)

variation $\swarrow$

$$\sum_{i=1}^n (y_i - \bar{y})^2 = S_{xy}^2 / S_{xx} + \left(S_{yy} - \frac{S_{xy}^2}{S_{xx}} \right)$$

$\Downarrow$

$$S_{yy} = \frac{S_{xy}^2}{S_{xx}} + \left(S_{yy} - \frac{S_{xy}^2}{S_{xx}} \right)$$

total variability of y $\nearrow$ variability explained by linear relation $\uparrow$ Residual or unexplained variability $\nwarrow$

For the straight line to provide a good fit to the data, the sum of squares due to regression $\frac{S_{xy}^2}{S_{xx}}$ should be a substantial proportion of the total sum of squares S_{yy} .

Sum of squares due to regression

Total sum of squares of y

$$\frac{\text{Sum of squares due to regression}}{\text{Total sum of squares of y}} = \frac{\frac{S_{xy}^2}{S_{xx}}}{S_{yy}} = \frac{S_{xy}^2}{S_{xx} S_{yy}} = r^2$$

(sample correlation coefficient)