#### Lecture 8: Model- and density-based clustering

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## Model-based clustering

In Quadratic Discriminant Analysis (QDA) we assumed

$$p(\mathbf{x}|i) = N(\mathbf{x}|\boldsymbol{\mu}_i, \boldsymbol{\Sigma}_i) \text{ and } p(i) = \pi_i$$

This is known as a Gaussian Mixture Model (GMM) for x where

$$p(\mathbf{x}) = \sum_{i=1}^{K} p(i)p(\mathbf{x}|i) = \sum_{i=1}^{K} \pi_i N\left(\mathbf{x}|\boldsymbol{\mu}_i, \boldsymbol{\Sigma}_i\right)$$

QDA used that the classes  $i_l$  and feature vectors  $\mathbf{x}_l$  of the observations were known to calculate  $\pi_i$ ,  $\boldsymbol{\mu}_i$  and  $\boldsymbol{\Sigma}_i$ .

What if we only know the features x<sub>l</sub>?

#### Maximum Likelihood for GMMs?

The log-likelihood for the data  $\mathbf{X} \in \mathbb{R}^{n \times p}$  and all unknowns

$$\boldsymbol{\theta} = (\pi_1, \boldsymbol{\mu}_1, \boldsymbol{\Sigma}_1, \dots, \pi_K, \boldsymbol{\mu}_K, \boldsymbol{\Sigma}_K)$$

is

$$\log(p(\mathbf{X}|\boldsymbol{\theta})) = \sum_{l=1}^{n} \log\left(\sum_{i=1}^{K} \pi_{i} N\left(\mathbf{x}_{l} | \boldsymbol{\mu}_{i}, \boldsymbol{\Sigma}_{i}\right)\right)$$

Taking the gradient (with chain-rule) and solving for some  $\mu_i$  gives

$$\boldsymbol{\mu}_{i} = \frac{\sum_{l=1}^{n} \eta_{li} \mathbf{x}_{l}}{\sum_{l=1}^{n} \eta_{li}} \quad \text{where} \quad \eta_{li} = \frac{\pi_{i} N(\mathbf{x}_{l} | \boldsymbol{\mu}_{i}, \boldsymbol{\Sigma}_{i})}{\sum_{j=1}^{K} \pi_{j} N(\mathbf{x}_{l} | \boldsymbol{\mu}_{j}, \boldsymbol{\Sigma}_{j})}$$

Note: There is a non-linear cyclic dependence between  $\eta_{li}$  and  $\mu_{i}$ .

#### **Expectation-Maximization for GMMs**

Finding the MLE for parameters  $\theta$  in GMMs results in an iterative process called **Expectation-Maximization (EM)** 

- 1. Initialize  $\theta$
- 2. E-Step: Update

$$\eta_{li} = \frac{\pi_i N(\mathbf{x}_l | \boldsymbol{\mu}_i, \boldsymbol{\Sigma}_i)}{\sum_{j=1}^K \pi_j N(\mathbf{x}_l | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)}$$

3. M-Step: Update

$$\boldsymbol{\mu}_{i} = \frac{\sum_{l=1}^{n} \eta_{li} \mathbf{x}_{l}}{\sum_{l=1}^{n} \eta_{li}} \qquad \boldsymbol{\pi}_{i} = \frac{\sum_{l=1}^{n} \eta_{li}}{n}$$
$$\boldsymbol{\Sigma}_{i} = \frac{1}{\sum_{l=1}^{n} \eta_{li}} \sum_{l=1}^{n} \eta_{li} (\mathbf{x}_{l} - \boldsymbol{\mu}_{i}) (\mathbf{x}_{l} - \boldsymbol{\mu}_{i})^{T}$$

4. Repeat steps 2 and 3 until convergence

#### GMM clustering example

- Yellow and green clusters share a covariance matrix
- The blue cluster has a different one
- GGM clustering on only the data points without knowledge of the class labels recovers the covariance structures and clusters



# Why does Expectation-Maximization work?

#### Likelihood of the complete data

- ▶ Assume that the classes  $i_l$  are known and code them as  $z_{lj} = 1$  if  $i_l = j$  and  $z_{lj} = 0$  otherwise. Collect them in  $\mathbf{Z} \in \mathbb{R}^{n \times K}$ .
- (X, Z) are called the complete data, and incomplete data when only X is observed
- The class assignments Z are called latent variables
- Complete data likelihood

$$\log(p(\mathbf{X}, \mathbf{Z} | \boldsymbol{\theta})) = \sum_{l=1}^{n} \sum_{i=1}^{K} z_{li} \left( \log(\pi_i) + \log(N(\mathbf{x}_l | \boldsymbol{\mu}_i, \boldsymbol{\Sigma}_i)) \right)$$

and the parameters in  $\theta$  are easy to estimate (QDA).

Incomplete data likelihood

$$\log(p(\mathbf{X}|\boldsymbol{\theta})) = \sum_{l=1}^{n} \log\left(\sum_{i=1}^{K} \pi_{i} N\left(\mathbf{x}_{l} | \boldsymbol{\mu}_{i}, \boldsymbol{\Sigma}_{i}\right)\right)$$

#### Decomposing the incomplete data likelihood

If we knew Z then

$$p(\mathbf{X}|\theta) = \frac{p(\mathbf{X}, \mathbf{Z}|\theta)}{p(\mathbf{Z}|\mathbf{X}, \theta)},$$
 i.e.

 $\log(p(\mathbf{X}|\boldsymbol{\theta})) = \log(p(\mathbf{X},\mathbf{Z}|\boldsymbol{\theta})) - \log(p(\mathbf{Z}|\mathbf{X},\boldsymbol{\theta}))$ 

a decomposition of the log-likelihood for X given θ
For any density q(Z) it holds that

$$\log(p(\mathbf{X}|\boldsymbol{\theta})) = \log\left(\frac{p(\mathbf{X}, \mathbf{Z}|\boldsymbol{\theta})}{q(\mathbf{Z})}\right) - \log\left(\frac{p(\mathbf{Z}|\mathbf{X}, \boldsymbol{\theta})}{q(\mathbf{Z})}\right)$$

• Average over  $\mathbf{Z}$  according to the density  $q(\mathbf{Z})$ 

$$\log(p(\mathbf{X}|\boldsymbol{\theta})) = \mathbb{E}_{q(\mathbf{Z})}\left[\log\frac{p(\mathbf{X},\mathbf{Z}|\boldsymbol{\theta})}{q(\mathbf{Z})}\right] - \mathbb{E}_{q(\mathbf{Z})}\left[\log\frac{p(\mathbf{Z}|\mathbf{X},\boldsymbol{\theta})}{q(\mathbf{Z})}\right]$$

It can be shown (using Jensen's inequality) that

$$\mathbb{E}_{q(\mathbf{Z})}\left[\log\frac{p(\mathbf{Z}|\mathbf{X},\boldsymbol{\theta})}{q(\mathbf{Z})}\right] \leq 0$$

with equality if  $q(\mathbf{Z}) = p(\mathbf{Z}|\mathbf{X}, \boldsymbol{\theta})$ 

#### **Expectation-Maximization**

#### New target function: Maximize

$$\log(p(\mathbf{X}|\boldsymbol{\theta})) = \mathbb{E}_{q(\mathbf{Z})} \left[ \log \frac{p(\mathbf{X}, \mathbf{Z}|\boldsymbol{\theta})}{q(\mathbf{Z})} \right] - \mathbb{E}_{q(\mathbf{Z})} \left[ \log \frac{p(\mathbf{Z}|\mathbf{X}, \boldsymbol{\theta})}{q(\mathbf{Z})} \right]$$
  
with respect to  $q(\mathbf{Z})$  and  $\boldsymbol{\theta}$ 

1. **Expectation step:** For given parameters  $\theta^{(m)}$  the density  $q(\mathbf{Z}) = p(\mathbf{Z}|\mathbf{X}, \theta^{(m)})$  minimizes the second term and thereby maximizes the first one. Set

$$Q(\boldsymbol{\theta}, \boldsymbol{\theta}^{(m)}) = \mathbb{E}_{p(\mathbf{Z}|\mathbf{X}, \boldsymbol{\theta}^{(m)})} \left[ \log \frac{p(\mathbf{X}, \mathbf{Z}|\boldsymbol{\theta})}{p(\mathbf{Z}|\mathbf{X}, \boldsymbol{\theta}^{(m)})} \right]$$

2. Maximization step: Maximize the first term with

$$\boldsymbol{\theta}^{(m+1)} = \operatorname*{arg\,max}_{\boldsymbol{\theta}} Q(\boldsymbol{\theta}, \boldsymbol{\theta}^{(m)})$$

The incomplete data likelihood increases in each step until convergence to a **local maximum**.

#### Applying EM to the GMM clustering problem

Given **X** and  $\theta^{(m)}$ 

$$p(\mathbf{Z}|\mathbf{X}, \boldsymbol{\theta}^{(m)}) = \frac{p(\mathbf{X}, \mathbf{Z}|\boldsymbol{\theta}^{(m)})}{p(\mathbf{X}|\boldsymbol{\theta}^{(m)})} = \frac{\prod_{l=1}^{n} \prod_{i=1}^{K} (\pi_{i} N(\mathbf{x}_{l}|\boldsymbol{\mu}_{i}, \boldsymbol{\Sigma}_{i}))^{z_{li}}}{\sum_{j=1}^{K} \pi_{j} N\left(\mathbf{x}|\boldsymbol{\mu}_{j}, \boldsymbol{\Sigma}_{j}\right)}$$

and it holds that

$$\mathbb{E}_{p(\mathbf{Z}|\mathbf{X},\boldsymbol{\theta}^{(m)})}[z_{li}] = \frac{\pi_i N(\mathbf{x}_l|\boldsymbol{\mu}_i,\boldsymbol{\Sigma}_i)}{\sum_{j=1}^K \pi_j N(\mathbf{x}_l|\boldsymbol{\mu}_j,\boldsymbol{\Sigma}_j)} = \boldsymbol{\eta}_{li}$$

the so-called **responsibility** of class *i* for having generated the observation  $\mathbf{x}_l$ .

This results in

$$Q(\boldsymbol{\theta}, \boldsymbol{\theta}^{(m)}) = \sum_{l=1}^{n} \sum_{i=1}^{K} \eta_{li} \left( \log(\pi_i) + \log(N(\mathbf{x}_l | \boldsymbol{\mu}_i, \boldsymbol{\Sigma}_i)) \right) \quad (+ \text{ const})$$

which is maximized by estimates seen earlier weighted by  $\eta_{li}$ .

A final clustering can be selected with

 $C(\mathbf{x}_l) = \arg\max_i \eta_{li}$ 

or responsibilities can be used as a soft clustering

**Cluster count selection:** Model selection criteria for MLE can be used, e.g. minimal **Bayesian Information Criterion (BIC)** 

BIC(K) = 
$$-2\log(p(\mathbf{X}|\boldsymbol{\theta}, K))$$
  
+  $\log(n) \cdot \underbrace{[(K-1) + K \cdot p + K \cdot \frac{p(p+1)}{2}]}_{\text{number of model parameters}}$ 

where n is much larger than the number of model parameters

- Centering one mixture component on an observation (i.e. µ<sub>i</sub> = x<sub>l</sub> for some *i* and *l*) and letting its variance go to zero can drive the likelihood to infinity
  - Outside of scope solution: Bayesian framework and Inverse-Wishart prior on Σ<sub>i</sub>
  - Initialize Σ<sub>i</sub> with large enough variances and potentially restart if bad convergence
- ▶ Like k-means, this algorithm is sensitive to starting values

#### GMMs and EM for classification

#### **GMM for classification**

# In QDA $p(\mathbf{x}|i) = N(\mathbf{x}|\boldsymbol{\mu}_i, \boldsymbol{\Sigma}_i)$ can only capture **elliptic class** shapes.

Assume features are described by a GMM, i.e.

$$p(\mathbf{x}|i) = \sum_{m=1}^{M_i} \pi_{im} N(\mathbf{x}|\boldsymbol{\mu}_{im}, \boldsymbol{\Sigma})$$

where

- *M<sub>i</sub>* components for class *i*
- $\pi_{im}$  is the probability of mixture component *m* for class *i*
- Covariance matrix Σ is assumed to be constant across mixture components and classes

**Component membership**  $z_{lm}$  is a latent variable for the observation  $(\mathbf{x}_l, i_l)$  with  $z_{lm} = 1$  if  $\mathbf{x}_l$  is in component  $m \in \{1, ..., M_{i_l}\}$  and  $z_{lm} = 0$  otherwise

#### **Mixture DA**

Finding the MLE for the mixture DA parameters can be achieved through **Expectation-Maximization (EM)** 

- 1. Initialize  $\theta$
- 2. E-Step: Update

$$\eta_{lm} = \frac{\pi_{i_lm} N(\mathbf{x}_l | \boldsymbol{\mu}_{i_lm}, \boldsymbol{\Sigma})}{\sum_{j=1}^{M_{i_l}} \pi_{i_lj} N(\mathbf{x}_l | \boldsymbol{\mu}_{i_lj}, \boldsymbol{\Sigma})}$$

3. M-Step: Update

$$\boldsymbol{\mu}_{im} = \frac{\sum_{i_l=i} \eta_{lm} \mathbf{x}_l}{\sum_{i_l=i} \eta_{lm}} \qquad \boldsymbol{\pi}_{im} = \frac{\sum_{i_l=i} \eta_{lm}}{n_i}$$
$$\boldsymbol{\Sigma} = \frac{1}{n} \sum_{i=1}^{K} \sum_{i_l=i} \sum_{m=1}^{M_i} \eta_{lm} (\mathbf{x}_l - \boldsymbol{\mu}_{im}) (\mathbf{x}_l - \boldsymbol{\mu}_{im})^T$$

4. Repeat steps 2 and 3 until convergence



## **Density-based clustering**

#### Yet another approach to clustering

- Most methods discussed so far have problems with odd, non-convex shapes
- What about noise? Some observations might not fit into any cluster
- New cluster definition: Clusters are dense regions in feature space
  - What is dense?
  - How to find groups and separate the noise?

Naive approach: Find points surrounded by many other points and connect them to a cluster. Points that do not end up in a cluster are noise.



#### Notation in density-based clustering

The presented methodology has **two tuning parameters**  $\varepsilon > 0$  and  $n_{\min} \in \mathbb{N}$ . Assume each observation is a **point** p **in a database/dataset** D and there is a **distance measure** d(p,q).

- ► *ε*-neighbourhood of *p*:  $N_{\varepsilon}(p) = \{q \in D | d(p,q) \le \varepsilon\}$
- ▶ **Core point:** A  $p \in D$  s.th.  $|N_{\varepsilon}(p)| \ge n_{\min}$
- ▶ p is directly density-reachable from a core-point q if p ∈ N<sub>ε</sub>(q)
- ▶ p is density-reachable from a core-point q if there is a chain q = p<sub>1</sub>, p<sub>2</sub>, ..., p<sub>m</sub> = p s.th. p<sub>i+1</sub> is directly density-reachable from p<sub>i</sub>
- *p* and *q* are **density-connected** if there is a core-point *o* s.th. *p* and *q* are density-reachable from *o*

A **cluster** *C* is a set of points in *D* s.th.

- 1. If  $p \in C$  and q is density-reachable from p then  $q \in C$  (maximality)
- 2. For all  $p, q \in C$ : p and q are density-connected (connectivity)
- This leads to three types of points
  - 1. Core points: Part of a cluster and at least  $n_{\min}$  points in neighbourhood
  - 2. Border points: Part of a cluster but not core points
  - 3. Noise: Not part of any cluster

Note: Border points can have non-unique cluster assignments

#### **Computational procedure:**

- 1. Go through each point p in the dataset D
- 2. If it has already been processed take the next one
- 3. Else determine its  $\varepsilon$ -neighbourhood. If less than  $n_{\min}$  points in neighbourhood, label as noise. Otherwise, start a new cluster.
- 4. Find all points that are density-reachable from p and add them to the cluster.

- Controls how easy it is to connect components in a cluster
  - Too small and most points are core points, creating many small clusters
  - Too large and few points are core points, leading to many noise labelled observations
- A cluster has by definition at least  $n_{\min}$  points
- Choice of  $n_{\min}$  is very dataset dependent
- Tricky in high-dimensional data (curse of dimensionality, everything is far apart)

#### Dependence on $\varepsilon$

- Controls how much of the data will be clustered
  - Too small and small gaps in clusters cannot be bridged, leading to isolated islands in the data
  - Too large and everything is connected
- Choice of ε is also dataset dependent but there is a decision tool
  - Determine distance to the k nearest neighbours for each point in the dataset
  - Inside clusters, increasing k should not lead to a large increase of d
  - The optimal e is supposed to be roughly at the knee



- DBSCAN is able to cluster points in the situations advertised and correctly identifies noise points
- Very sensitive to the choice of tuning parameters



- Expectation-Maximization allows maximum likelihood estimation even in situation where additional data would be necessary
- Both clustering and classification methods profit from using Gaussian Mixture Models
- Density-based clustering allows to capture complex shapes and the identification of noise during clustering