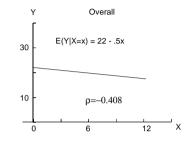
Effect reversal

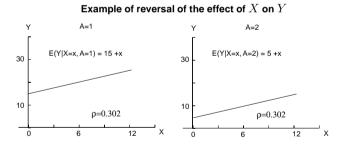
or: why it is important to study relations among explanatory variables

- Example for variables Y, X, A
- Example for variables A, B, C
- Example for variables Y, X, Z
- A general condition when it cannot occur

As a consequence of strongly associated explanatory variables X, A a reversal in the effect of X on Y occurs, when A is ignored



Slide 1



Here X is strongly dependent on A since

 $E(X \mid A = 1) = 5 \text{ and } E(X \mid A = 2) = 11$

and has little variability given $A: \mathrm{var}(X \mid A=1) = \mathrm{var}(X \mid A=2) = 1$



Details to the example

| Levels | Me | ans | Vari | ances | | |
|---------------------|----|-----|------|-------|----------------|--------|
| of \boldsymbol{A} | Y | X | Y | X | $\sigma_{X,Y}$ | P(A=i) |
| i = 1 | 20 | 5 | 11 | 1 | 1 | .5 |
| i = 2 | 16 | 11 | 11 | 1 | 1 | .5 |
| Overall | 18 | 8 | 15 | 10 | -5 | |

Slide 3

with e.g.

$$\sigma_{XY} = E_A \{ \operatorname{cov}(Y, X \mid A) \} + \operatorname{cov}_A \{ E(Y \mid A), E(X \mid A) \}$$

= (.5 × 1 + .5 × 1)
+ {.5 × (20 - 18)(5 - 8) + .5 × (16 - 18)(11 - 8)}
= -5

Slide 4

In two clinics a new treatment (j = 1) is better, but....

| | <i>k</i> = | - , | Clinic k = | inic $k=2$ | | | $\begin{array}{c} \text{counts for } AB \\ \text{ignoring } C \end{array}$ | | |
|------------------|--------------|-------|---------------|--------------|---|--------------|--|--|--|
| A, Treat- | B, Treatment | | B, Treatment | | | B, Treatment | | | |
| ment | j = 1 | j = 2 | j = 1 | j = 2 | | j = 1 | j = 2 | | |
| success | new | conv. | new | conv. | _ | new | conv. | | |
| $i=1: {\sf yes}$ | 60 | 4 | 30 | 200 | | 90 | 204 | | |
| $(P_{1 jk})$ | (30%) | (20%) | (75%) | (50%) | | (38%) | (49%) | | |
| i=2: no | 140 | 16 | 10 | 200 | _ | 150 | 216 | | |
| sum | 200 | 20 | 40 | 400 | _ | 240 | 420 | | |
| rel. chance | 30/20 = 1, 5 | | 75/50 | 75/50 = 1, 5 | | | 38/49 = 0,78 | | |

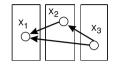
Slide 5

| Replicated relative chances are preserved sin | Ince D | and C | are not assoc | lated |
|---|--------|-------|---------------|-------|
|---|--------|-------|---------------|-------|

| | | со | counts from | | | | | |
|------------------------------------|------------------|--------------------|-------------------|---------------------|------|----------------|---------------------|--|
| | k = 1 | | <i>k</i> = | k = 2 | | both clinics | | |
| A, Treat- | B, Treatment | | B, Tre | atment | В, | B, Treatment | | |
| ment | j = 1 | j = 2 | j = 1 | j = 2 | j = | 1 | j = 2 | |
| success | new | conv. | new | conv. | ne | W | conv. | |
| $i=1:$ yes $(P_{1 jk})$ i=2: no | 6 (30%) 14 | 40 (20%) 160 | 30 (75%) 10 | 200 (50%) 200 | (60% | 36 %) 24 | 240 (40%) 360 | |
| sum | 20 | 200 | 40 | 400 | 6 | 60 | 600 | |
| rel. chance | 30/20 = 1,5 | | 75/50 | = 1, 5 | 60/ | 60/40 = 1,5 | | |

Overall effect in a Gaussian linear system

For mean-centered variables X_1, X_2, X_3 the **triangular system** with the graph



corresponding linear equations can be written as:

$$E(X_1 \mid X_2, X_3) = \beta_{1|2.3}X_2 + \beta_{1|3.2}X_3$$
$$E(X_2 \mid X_3) = \beta_{2|3}X_3$$
$$E(X_3) = 0$$

Slide 7

The ${\it overall}\ {\it effect}$ of X_3 on X_1 is the sum of effects of 'two paths' since

$$\begin{split} \mathbf{E}(X_1 \mid X_3) &= \mathbf{E}_{X_2 \mid X_3} \mathbf{E}(X_1 \mid X_2, X_3) \\ &= \mathbf{E}_{X_2 \mid X_3} (\beta_{1 \mid 2.3} X_2 + \beta_{1 \mid 3.2} X_3) \\ &= \beta_{1 \mid 2.3} \mathbf{E}_{X_2 \mid X_3} (X_2) + \beta_{1 \mid 3.2} X_3 \\ &= \beta_{1 \mid 2.3} \beta_{2 \mid 3} X_3 + \beta_{1 \mid 3.2} X_3 \\ &\text{or} \\ &\beta_{1 \mid 3} X_3 &= (\beta_{1 \mid 2.3} \beta_{2 \mid 3} + \beta_{1 \mid 3.2}) X_3 \end{split}$$

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Suppose a R esponse , T reatment, and B ackround variable are standardized to have mean zero and variance 1; they be linearly related and correlated as

$$\operatorname{cor} = \left(\begin{array}{ccc} 1 & -.32 & .544 \\ . & 1 & -.8 \\ . & . & 1 \end{array} \right)$$

then - by computing least squares regression coefficents - we get

$$\begin{split} E_{lin}(R \mid T,B) &= .32 \times T + .8 \times B \\ E_{lin}(T \mid B) &= -.8 \times B \end{split}$$

and $\beta_{\rm RT}=-.32$, i.e. effect reversal

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But for the same $E_{lin}(R \mid T, B)$ with T, B uncorrelated

 $\begin{array}{lll} E_{lin}(R \mid T,B) &=& .32 \times T + .8 \times B \\ \\ E_{lin}(T \mid B) &=& 0 \times B \end{array}$

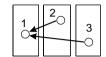
 $eta_{
m RT}=.32$, i.e. effect is preserved

The correlation matrix is

$$cor = \left(\begin{array}{ccc} 1 & .32 & .8 \\ . & 1 & 0 \\ . & . & 1 \end{array} \right)$$

On the overall effect for general distributions

generated over the graph



The joint density is in condensed notation

$$f_{123} = f_{1|23} f_2 f_3$$

The overall effect of variable 3 on variable 1 is the dependence in

$$f_{1|2} = \int f_{1|23} f_3 dx_3$$

Slide 11

 $X_2 \perp \!\!\!\perp X_3$ is sufficient for no effect reversal in the case variable X_1 depends monotonically on X_2 for all (possible reordered) levels of X_3 (Cox and Wermuth, 2003)

This means that at least qualitatively one obtains the same conclusions regarding the direction of dependence of variable 1 on variable 2, no matter whether variable 3 is explicitly considered (with $f_{1\mid 23}$) or not (with $f_{1\mid 2}$)

Without monotonous dependence - as e.g. in the Lienert data - no such conclusions are possible

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